# Lecture 2: Distinct element counting 

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NOTE: THESE NOTES HAVE NOT BEEN EDITED OR CHECKED FOR CORRECTNESS

## 1 Overview

In the last lecture we looked at examples of property testing, streaming, and testing distributions. In this lecture we cover estimation of a Bernoulli random variable and the number of distinct elements in a stream.

## 2 Bernoulli random variables

Given a weighted coin (comes up heads with probability $p \in[0,1]$ ), how many flips do we need to estimate $p$ ? Lets assume we take $k$ samples. We will call the sample mean $\hat{p}$. As $n \rightarrow \infty, \hat{p} \rightarrow p$, meaning that this is a reasonable way to guess the value of $p$. In order to figure out how big we need to make $n$ for our answer to be reasonable, it would be useful to find the variance of $\hat{p}$.

$$
\begin{aligned}
\operatorname{Var}(\hat{p}) & =\operatorname{Var}\left(\frac{1}{n^{2}} \sum x_{i}\right) \\
& =\frac{1}{n^{2}} \operatorname{Var}\left(\sum x_{i}\right) \\
& =\frac{1}{n} \operatorname{Var}\left(x_{1}\right) \\
& =\frac{1}{n} \mathbb{E}\left[\left(x_{1}-\mathbb{E}\left[x_{1}\right]\right)^{2}\right] \\
& =\frac{1}{n}\left(p(1-p)^{2}+(1-p) p^{2}\right) \\
& =\frac{1}{n}(p(1-p)) \\
& \leq \frac{1}{4 n}
\end{aligned}
$$

This tells us that the standard deviation of $\hat{p}$ is at most $\sqrt{n p}$. Thus we can expect that with n samples we will get that $\hat{p} \in[p-\sqrt{p / n}, p+\sqrt{p / n}]$.

Suppose you want to estimate $p$ within an additive error of $\epsilon$ while having a failure probability of at most $\frac{1}{4}$. What $n$ value would we need? By the above reasoning, we get that setting $n=O\left(\frac{1}{\epsilon^{2}}\right)$ gives us that $\hat{p} \in[p-O(\epsilon), p+O(\epsilon)]$ with sufficient probability (by Chebyshev's inequality).

## 3 Basic probability inequalities

There are two simple inequalities that show up a lot when dealing with probabilities. Markov's inequality gives us that for any non-negative random variable $x, \mathbb{P}[x \geq t] \leq \frac{\mathbb{E}[x]}{t}$. Chebyshev's inequality gives us that for any integrable random variable with finite mean $\mu$ and standard deviation $\sigma, \mathbb{P}[|x-\mu| \geq t \sigma] \leq \frac{1}{t^{2}}$. Note that this tells us that the probability that a random variable is more than 2 standard deviations from the mean is less than $\frac{1}{4}$.

## 4 Mean estimation

Suppose I have an unknown distribution $D$ with an unknown mean $\mu$ whose standard deviation is at most $\sigma$. How many samples will I need from $D$ to estimate $\mu$ to within an additive factor of $\epsilon \sigma$ with $\frac{3}{4}$ probability?

We will take the empirical mean as the variable $\hat{\mu}$. We know that $\operatorname{Var}(\hat{\mu}) \leq \sigma^{2} / n$, where $n$ is the number of samples we take. Thus setting $n=\frac{4}{\epsilon^{2}}$ gives us that the standard deviation of $\hat{\mu}$ is at most $\frac{\epsilon \sigma}{2}$. By Chebyshev's inequality, this implies that $\hat{\mu} \in[\mu-\epsilon \sigma, \mu+\epsilon \sigma]$ with a probability of at least $\frac{3}{4}$.

## 5 Streaming distinct elements

Let's say you have a stream of items that you only get to pass through once. Your goal is to estimate the number of distinct elements ( $n$ ) in the stream. What is the least amount of information you need to store to get a good estimate of the number of distinct elements in the stream? Lets start with a simpler problem: we are promised that either $n<T$ or $n>2 T$ and we want to tell which is true.

To solve this we will first construct a random hash function $h: U \rightarrow[T]$. We will then pick $k$ elements of $T$ and for each selected element of $T$, we will increment a counter the first time that a hash hits it. We know that each element of $T$ has a probability of $\left(1-\frac{1}{T}\right)^{n} \approx e^{-n / T}$ to get an element of the stream to hash to it. When $n<T$ this gives us a probability of at most .63 and when $n>2 T$ the probability is at least .86 . Thus our problem gets reduced to selecting a large enough $k$ (here, representing the number of parallel runs) such that we can distinguish between these two Bernoulli random variables. From before, we know that $O\left(\frac{1}{\epsilon^{2}}\right)$ space should be enough to distinguish between $(1-\epsilon) T$ and $(1+\epsilon) T$ unique elements.

We now move to the original question of estimating number of unique elements. One idea is to repeat the above algorithm in parallel for different values of $T$. In order to get ( $1 \pm \epsilon$ ) error guarantee, we can repeat the above algorithm in parallel for $T=1,(1+\epsilon),(1+\epsilon)^{2}, \cdots,(1+\epsilon)^{\log _{1+\epsilon}(N)}$. For most of the $T$ values, the algorithm will say either less than or greater than. But a few of the runs in middle will be confused as they don't satisfy $<T$ or $>2 T$ condition.
A solution to deal with this is to repeat many times. We saw earlier in the coin flip example that to get constant success probability, $O\left(1 / \epsilon^{2}\right)$ flips are required. We'll see later in the course it's possible to get $\geq 1-\delta$ success probability by tossing $O\left(\frac{1}{\epsilon^{2}} \log \left(\frac{1}{\delta}\right)\right)$ flips. Therefore, using this information,
and doing union bound over constant failure probability of all the runs, the space needed for is:

$$
\begin{aligned}
& O(\underbrace{\log _{1+\epsilon}(N)}_{\text {initial number of runs proposed }} \cdot \underbrace{\frac{1}{\epsilon^{2}}}_{\text {space complexity for each run }} \cdot \underbrace{\log (\log 1+\epsilon(N))}_{\text {more runs for union-bounding failure prob }} \\
& =O\left(\frac{1}{\epsilon} \log (N) \cdot \frac{1}{\epsilon^{2}} \cdot \log \left(\frac{1}{\epsilon} \log (N)\right)\right)
\end{aligned}
$$

In order to get success probability $\geq 1-\delta$, the space required is:

$$
O\left(\frac{1}{\epsilon} \log (N) \cdot \frac{1}{\epsilon^{2}} \cdot \log \left(\frac{1}{\delta \epsilon} \log (N)\right)\right)
$$

We now look at another simpler algorithm. We'll hash the incoming units to a value between $[0,1]$. That is, we consider a hash function $h: U \rightarrow[0,1]$. Let's analyze the expected minimum value that any unit is hashed to. Let $S$ be the set comprising all elements we see. Let $y=\min _{X \in S} h(X)$, then

$$
\begin{aligned}
\mathbb{E}[y] & =\frac{1}{n+1} \\
\text { Proof: } \mathbb{P}[y \geq 1-t] & =t^{n} \\
\Longrightarrow \mathbb{P}[y=1-t] & =n t^{n-1} \\
\mathbb{E}[1-y] & =\int_{t=0}^{1}\left(n t^{n-1}\right) t d t \\
& =\frac{n}{n+1} \\
\Longrightarrow \mathbb{E}[y] & =\frac{1}{n+1}
\end{aligned}
$$

We analyze the variance of $y$

$$
\begin{aligned}
\operatorname{Var}(y) & =\operatorname{Var}(1-y) \\
& =\mathbb{E}\left[(1-y)^{2}\right]-(\mathbb{E}[1-y])^{2} \\
\mathbb{E}\left[\left(1-y^{2}\right)\right] & =\int_{t=0}^{1}\left(n t^{n-1}\right) t^{2} d t=\frac{n}{n+2} \\
\Longrightarrow \operatorname{Var}(y) & =\frac{n}{n+2}-\left(\frac{n}{n+1}\right)^{2} \\
& =\frac{n}{(n+2)^{2}(n+1)} \\
& \approx \frac{1}{(n+1)^{2}}
\end{aligned}
$$

Algorithm: Hash the incoming units to a value between $[0,1]$. Let $y=\min _{X \in S} h(X)$. Output $\frac{1}{y}-1$. The above algorithm won't work well because Variance of $y$ is low. Basically, the minimum hashed value has good variation in it. Therefore, we'll use the same old technique of repeating experiment many times and taking the average

Algorithm: Hash the incoming units to a value between $[0,1]$ on $r$ different hash functions. let $y_{i}=\min _{X \in S} h_{i}(X)$. Output $\frac{1}{\frac{1}{r} \sum_{i=1}^{r} y_{i}}-1$.
Note that ( $1 \pm \epsilon$ ) factor approximation to $\frac{1}{r} \sum_{i=1}^{r} y_{i}$ will translate to ( $1 \pm \epsilon$ ) factor approximation for $\frac{1}{\frac{1}{r} \sum_{i=1}^{r} y_{i}}$ too. Therefore, we focus on getting

$$
\begin{array}{r}
\frac{1}{r} \sum_{i=1}^{r} y_{i} \in(1 \pm \epsilon) \mathbb{E}\left[\frac{1}{r} \sum_{i=1}^{r} y_{i}\right] \\
\mathbb{E}\left[\frac{1}{r} \sum_{i=1}^{r} y_{i}\right]=\mathbb{E}[y]=\frac{1}{n+1} \\
\Longrightarrow \frac{1}{r} \sum_{i=1}^{r} y_{i} \in \frac{1+\epsilon}{n+1} \\
\Longrightarrow\left|\frac{1}{r} \sum_{i=1}^{r} y_{i}-\frac{1}{n+1}\right| \leq \frac{\epsilon}{n+1}=\epsilon \sigma
\end{array}
$$

where $\sigma$ is the std. deviation as we calculated above. Using Chebyshev's inequality, we get that the number of samples $r$ needed is $r=O\left(\frac{1}{\epsilon^{2}}\right)$

Space Complexity: Need $O\left(1 / \epsilon^{2}\right)$ hash functions, to estimate each $y_{i}$. The space required to store $y_{i}$ depends on the resolution. A resolution of $1 / n^{2}$ is sufficient for our purposes, but we can do better. Since we only need $y_{i}$ to a $(1 \pm \epsilon)$ multiplicative factor, we can round our $y_{i}$ to the nearest $(1+\epsilon)^{-i}$ for $i \in \mathbb{Z}$. This gives $\frac{1}{\epsilon} \log (n)$ possible values for $y_{i}$, which only requires $\log \left(\frac{1}{\epsilon} \log (n)\right)$ bits of storage. This gives us a final space complexity of $O\left(\frac{1}{\epsilon^{2}} \log \left(\frac{1}{\epsilon} \log (n)\right)\right)$.

