

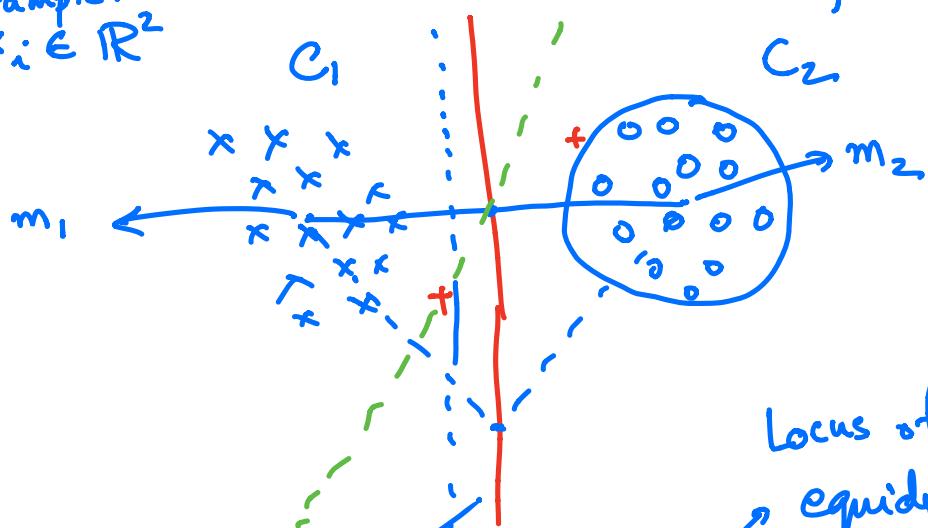
Feb 21, 2020

Classification

Training set: $(x_1, y_1), (x_2, y_2) \dots (x_n, y_n)$

Goal: Learn f that predicts class

Example: label of new (test) point x .
 $x_i \in \mathbb{R}^2$



$$\|x - m_1\|_2^2 = \|x - m_2\|_2^2 \quad \begin{matrix} \text{Locus of points} \\ \text{equidistant} \\ \text{to } m_1 \& m_2 \end{matrix}$$

$$x^T x - 2x^T m_1 + \|m_1\|_2^2 = x^T x - 2x^T m_2 + \|m_2\|_2^2$$

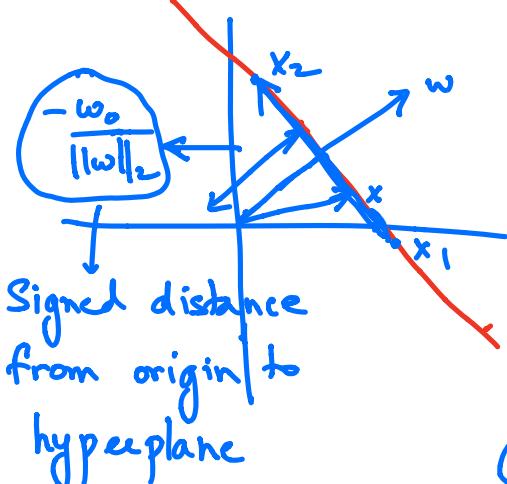
$$y(x) = (m_2 - m_1)^T x + \frac{1}{2} (\|m_1\|_2^2 - \|m_2\|_2^2) = 0$$

$$y(x) > 0, x \in C_2$$

$$y(x) < 0, x \in C_1$$

↓
 Example of a hyperplane (linear variety of
 $(d-1)$ dimensions when
 d is dimensionality
 of data)

$$H: w^T x + w_0 = 0$$



- Hyperplane

If x lies on the hyperplane,

$$w^T x + w_0 = 0$$

$$\left(\frac{w}{\|w\|_2}\right)^T x = -\frac{w_0}{\|w\|_2}$$

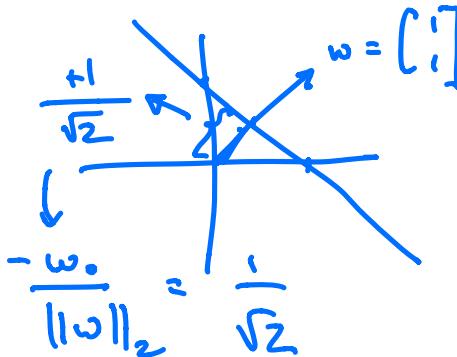
$$w^T x_1 + w_0 = 0$$

$$w^T x_2 + w_0 = 0$$

$$(outward) \quad w^T(x_1 - x_2) = 0$$

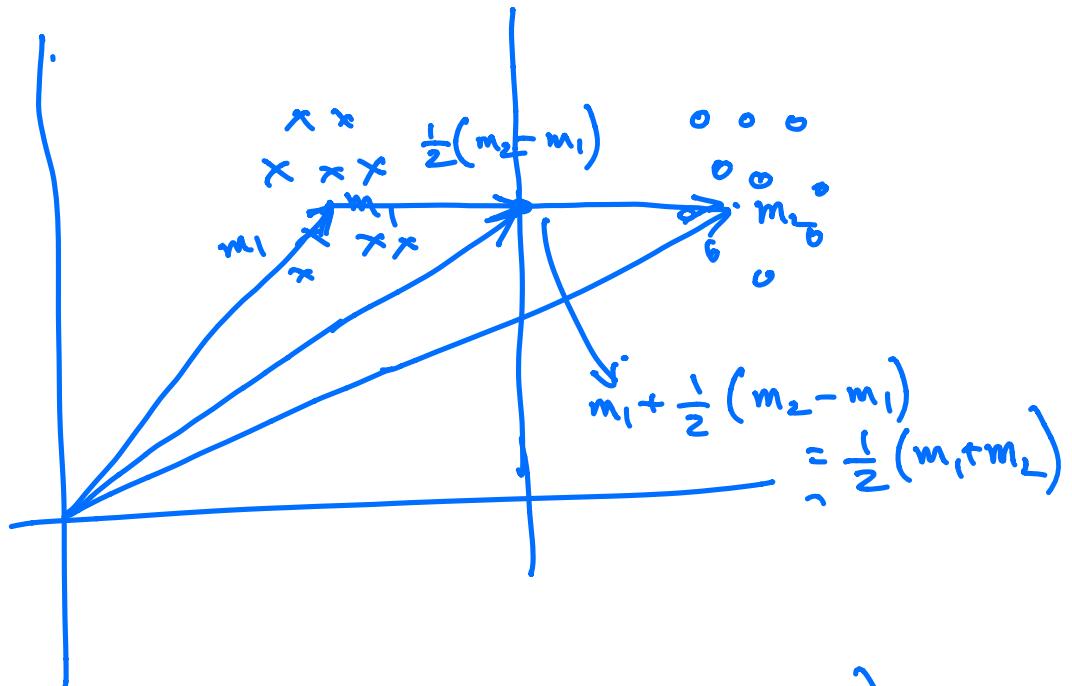
w is normal to (the points on)
 hyperplane

Example : $x_1 + x_2 = 1$



$$\begin{bmatrix} 1 & 1 \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} - 1 = 0$$

$$w = \begin{bmatrix} 1 \\ 1 \end{bmatrix}, \quad w_0 = -1$$



$$y(x) = (m_2 - m_1)^T x + \frac{1}{2} (m_1^T m_1 - m_2^T m_2)$$

$p(C_1|x)$ $\xrightarrow{\text{data likelihood}}$ $p(C_2|x)$ $\xrightarrow{\text{prior}}$
 $p(C_i|x)$ $\xleftarrow{\text{posterior}}$
$$p(x|C_i)p(C_i) \quad \quad \quad p(x)$$

$$p(C_2|x) = \frac{p(x|C_2)p(C_2)}{p(x)}$$

MAP rule

↓
Maximum a posteriori probability : $\arg \max_i p(C_i|x)$

Gaussian Model: $p(x|C_i) = \frac{1}{(2\pi)^{d/2} |\Sigma_i|^{1/2}} e^{-\frac{1}{2} (x-\mu_i)^T \Sigma_i^{-1} (x-\mu_i)}$

$$\underset{i}{\operatorname{argmax}} \log p(c_i|x) = \underset{i}{\operatorname{argmax}} \log(p(x|c_i)p(c_i))$$

$$\underset{i}{\operatorname{argmax}} \boxed{\log p(x|c_i) + \log p(c_i)}$$

$$\boxed{\log p(x|c_i)} = -\frac{d}{2} \log 2\pi - \frac{1}{2} \log |\Sigma_i| - \frac{1}{2} (x - \mu_i)^T \Sigma_i^{-1} (x - \mu_i)$$

$$\text{Decision Surface: } -\log p(c_1|x) = -\log p(c_2|x)$$

$$\boxed{\frac{d}{2} \log 2\pi + \frac{1}{2} \log |\Sigma_1| + \frac{1}{2} (x - \mu_1)^T \Sigma_1^{-1} (x - \mu_1) - \log p(c_1) = \frac{d}{2} \log 2\pi + \frac{1}{2} \log |\Sigma_2| + \frac{1}{2} (x - \mu_2)^T \Sigma_2^{-1} (x - \mu_2) - \log p(c_2)}$$

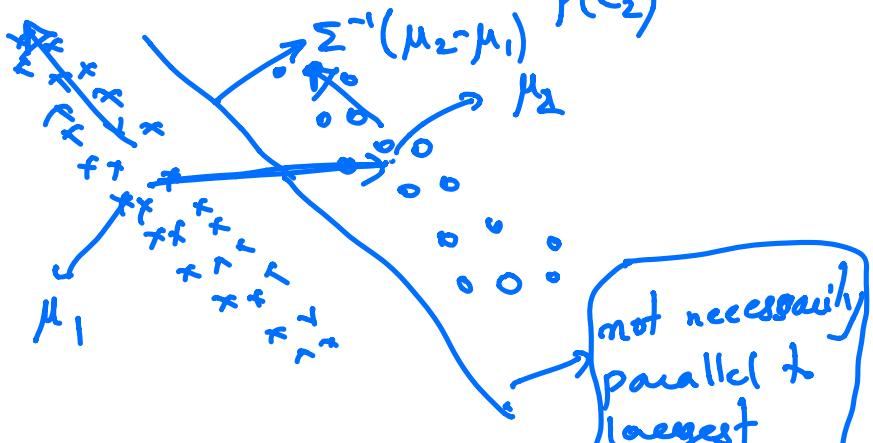
Case I: $\Sigma_1 = \Sigma_2 = I$

$$\frac{1}{2} (x - \mu_1)^T (x - \mu_1) - \log p(c_1) = \frac{1}{2} (x - \mu_2)^T (x - \mu_2) - \log p(c_2)$$

$$\frac{1}{2} (x^T x - 2\mu_1^T x + \|\mu_1\|^2) - \log p(c_1) =$$

$$\frac{1}{2} (x^T x - 2\mu_2^T x + \|\mu_2\|^2) - \log p(c_2)$$

$$(\mu_2 - \mu_1)^T x + \frac{1}{2} (\|\mu_1\|^2 - \|\mu_2\|^2) - \log \frac{p(c_1)}{p(c_2)} = 0$$



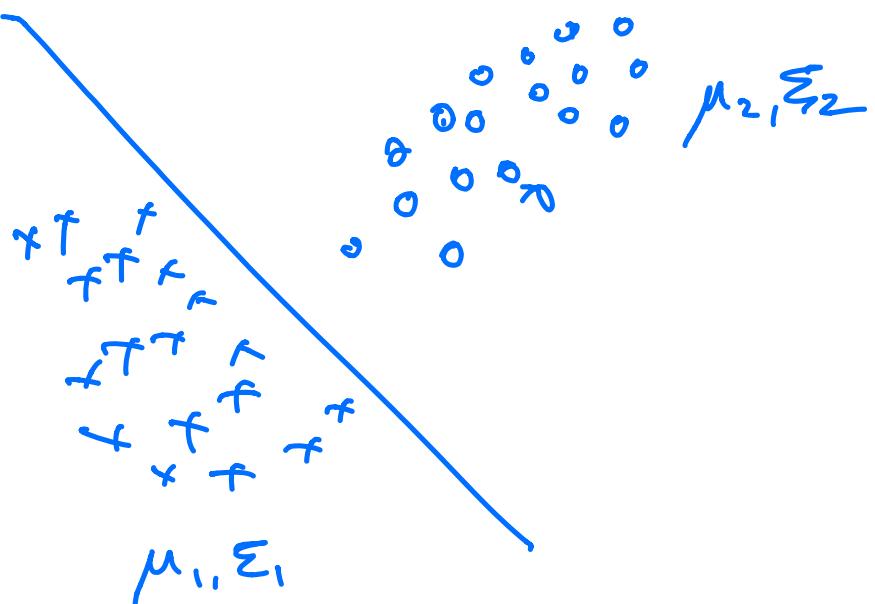
Case II: $\Sigma_1 = \Sigma_2 = \Sigma$ eigenvector
of Σ

$$\frac{1}{2}(x - \mu_1)^\top \Sigma^{-1} (x - \mu_1) - \log p(c_1) = \frac{1}{2}(x - \mu_2)^\top \Sigma^{-1} (x - \mu_2) - \log p(c_2)$$

Simplify: $(\mu_2 - \mu_1)^\top \Sigma^{-1} x + \frac{1}{2}(\mu_1 + \mu_2)^\top \Sigma^{-1} (\mu_1 - \mu_2)$
 $- \log \frac{p(c_1)}{p(c_2)} = 0$

Decision Surface is Linear

$$w^\top x + w_0 = 0$$
 $w = \Sigma^{-1}(\mu_2 - \mu_1)$



Classification : Regression Approaches

(x_i, y_i) , $x_i \in \mathbb{R}^d$, $y_i \in \mathbb{R}^k$ k-th position
 $y_i = [0, 0, \dots, \overset{\uparrow}{1}, 0, \dots, 0]$
 if $x_i \in C_k$

$$w_0 + w_1 x(1) + w_2 x(2) + \dots + w_d x(d) \rightarrow \text{Linear Fit}$$

Linear Discriminants: $w_k^T x + w_k$ for k-th class

$$\begin{array}{c}
 \text{d+1} \\
 \left[\begin{array}{cccccc}
 c_1 & \left\{ \begin{array}{cccccc}
 1 & x_i(1) & x_i(2) & \dots & x_i(d) \\
 \vdots & \vdots & \vdots & \ddots & \vdots \\
 c_2 & \vdots & \vdots & \vdots & \vdots \\
 \vdots & \vdots & \vdots & \vdots & \vdots \\
 c_k & \left\{ \begin{array}{cccccc}
 1 & \bar{x}_N(1) & \bar{x}_N(2) & \dots & \bar{x}_N(d)
 \end{array} \right. & \right. & \right. & \right. & \right. & \right. \\
 \end{array} \right]_{d+1} \left[\begin{array}{cccccc}
 w_{10} & w_{20} & \dots & w_{k0} \\
 \vdots & \vdots & \ddots & \vdots \\
 w_1 & w_2 & \dots & w_k
 \end{array} \right] \xrightarrow{\sim} \left[\begin{array}{cccccc}
 1 & 0 & \dots & y_1^T & 0 \\
 \vdots & \vdots & \ddots & y_2^T & 0 \\
 1 & 0 & \dots & \vdots & \vdots \\
 \vdots & \vdots & \vdots & \vdots & \vdots \\
 c_2 & 0 & 1 & \vdots & \vdots & 0 \\
 \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\
 c_k & 0 & 1 & \vdots & \vdots & 0 \\
 \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\
 y_N^T & & & & & 0
 \end{array} \right]
 \end{array}$$

$$XW \approx Y \Rightarrow \min_W \frac{1}{2} \|XW - Y\|_F^2$$

$$= \min_W \frac{1}{2} \text{Tr} \left[(XW - Y)^T (XW - Y) \right]$$

$$(X^T X) W^* = X^T Y \Rightarrow W^* = (X^T X)^{-1} X^T Y$$

xw^* is prediction on training data

$$x_i^T w \text{ could be } [-5 \quad -3 \quad 0 \quad 3 \quad 2 \quad 1 \quad 0.5 \quad 0.1]$$

and this is not a good approximation to

1-hot vectors

Least squares fit has obvious drawbacks

Special case of least squares fit:

2-class problem

$$y_i = \frac{N}{n_1} \text{ for each } x_i \in C_1, n_1 = |C_1|$$

$$y_i = -\frac{N}{n_2} \text{ for each } x_i \in C_2$$

Turns out to be equivalent to:

Fisher's Linear Discriminant

Logistic Regression

We had modeled each class as a Gaussian with covariance Σ :

$$\log \frac{p(C_i|x)}{p(C_j|x)} = \underbrace{\log \frac{p(C_i)}{p(C_j)}}_{w_0 + w^T x} - \frac{1}{2} (m_+ + m_-)^T \Sigma^{-1} (m_+ - m_-) + x^T \Sigma^{-1} (m_+ - m_-)$$

K-class problem

$$\log \frac{p(c_1|x)}{p(c_K|x)} = w_{10} + w_1^T x \quad -\textcircled{1}$$

$$\log p(c_2|x) = w_{20} + w_2^T x \quad -\textcircled{2}$$

:

$$\log \frac{p(c_{K-1}|x)}{p(c_K|x)} = w_{(K-1)0} + w_{K-1}^T x \quad -\textcircled{k-1}$$

$$\text{Write } p_i = p(c_i|x)$$

Add $\textcircled{1}, \textcircled{2}, \dots, \textcircled{k-1}$

$$e^{\textcircled{0}} \Rightarrow \frac{p_1}{p_K} = e^{w_{10} + w_1^T x}$$

$$\frac{p_1}{p_K} + \frac{p_2}{p_K} + \dots + \frac{p_{K-1}}{p_K} = \sum_{i=1}^{K-1} e^{w_i^T x}$$

$$1 - p_K = p_K \sum_{i=1}^{K-1} e^{w_i^T x}$$

$$\Rightarrow p_K = \frac{1}{1 + \sum_{i=1}^{K-1} e^{w_i^T x}}$$

$$p_i = \frac{e^{w_i^T x}}{1 + \sum_{i=1}^{K-1} e^{w_i^T x}}, \quad i=1, 2, \dots, K-1$$

$$\sigma(z) = \frac{1}{1+e^{-z}}$$

is called the logistic
Sigmoid function

$$\begin{aligned}\sigma'(z) &= +\frac{1 \cdot e^{-z}}{(1+e^{-z})^2} \\ &= \frac{e^{-z}}{1+e^{-z}} \cdot \frac{1}{1+e^{-z}} \\ &= (1-\sigma(z)) \cdot \sigma(z)\end{aligned}$$

$$\sigma' = \sigma(1-\sigma)$$

Parameters in logistic regression $\{w_{0i}, w_i\}_{i=0}^{k-1}$
 usually fit by maximum likelihood using the
 conditional likelihood given X .

Consider 2-class problem

$$\rightarrow p(C_1|x) = \boxed{p(w) = p}$$

$$\rightarrow p(C_2|x) = 1-p$$

(x_i, y_i) is training data
 $i=1, \dots, N$

Data Likelihood

$$\prod_{i=1}^N p^{y_i} (1-p)^{1-y_i}$$

$p(C_1|x)$ when $x \in C_1$

let $y_i = 1$ when $x_i \in C_1$

$y_i = 0$ when $x_i \in C_2$

$$\text{Max log-likelihood}$$

$$\max_w \lambda(w) = \max_w \left[\sum_{i=1}^N y_i \log p + (1-y_i) \log(1-p) \right]$$

$$1-p = \frac{1}{1+e^{w^T x_i}}, p = \frac{e^{w^T x_i}}{1+e^{w^T x_i}} \Rightarrow \log p = w^T x_i - \log(1+e^{w^T x_i})$$

$\nabla_w \ell(w) = 0$ - Minimizer will satisfy

$$\nabla_w \log p = x_i - \frac{1}{1+e^{w^T x_i}} \cdot e^{w^T x_i} \cdot x_i$$

$$\nabla_w \log(1-p) = -\frac{e^{w^T x_i}}{1+e^{w^T x_i}} \cdot x_i$$

$$\begin{aligned} \nabla_w \ell(w) &= \sum_{i=1}^N \left[y_i x_i \left(1 - \frac{e^{w^T x_i}}{1+e^{w^T x_i}} \right) + (1-y_i) x_i \left(-\frac{e^{w^T x_i}}{1+e^{w^T x_i}} \right) \right] \\ &= \sum x_i \left[y_i - \frac{y_i e^{w^T x_i}}{1+e^{w^T x_i}} - \frac{e^{w^T x_i}}{1+e^{w^T x_i}} + \frac{y_i e^{w^T x_i}}{1+e^{w^T x_i}} \right] \\ &= \boxed{\sum_{i=1}^N x_i \left[y_i - \frac{e^{w^T x_i}}{1+e^{w^T x_i}} \right] = 0} \end{aligned}$$

logistic regression
parameters satisfy this

$p(C_1 | x_i)$

$$= \sum_{i=1}^n x_i [y_i - (1 - \sigma(w^T x_i))]$$

$d+1$ non-linear equations in w
 ↓
 Re-weighted $(d+1)$ parameters
 Iterated Least squares method (IRLS)

Gradient Descent/Ascent

Newton's Method

Drawback is that each step
 of Gradient Descent is $O(N)$

Stochastic Gradient Descent

Regularization

$$\lambda \|w\|_2^2$$

$$\lambda \|w\|_1$$