Beyond the graphical Lasso: Structure learning via inverse covariance estimation

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Joint work with Martin Wainwright (UC Berkeley) & Peter Bühlmann (ETH Zürich)



- 2 Generalized inverse covariances
- 3 Linear structural equation models

4 Corrupted data



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Undirected graphical models

- Undirected graph G = (V, E)
- Joint distribution of (X_1, \ldots, X_p) , where |V| = p



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• Markov property:

$$(s,t) \notin E \implies X_s \perp \!\!\!\perp X_t \mid X_{\setminus \{s,t\}}$$

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• More generally, $X_A \perp\!\!\!\perp X_B \mid X_S$ when $S \subseteq V$ separates A from B

Directed graphical models

• Directed acyclic graph G = (V, E)



• Markov property:

$$X_j \perp \!\!\!\perp X_{\operatorname{Nondesc}(j)} \mid X_{\operatorname{Pa}(j)}, \qquad \forall j$$

• **Goal:** Edge recovery from *n* samples: $\{(X_1^{(i)}, X_2^{(i)}, ..., X_p^{(i)})\}_{i=1}^n$

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- High-dimensional setting: $p \gg n$, assume deg(G) $\leq d$
- Sources of corruption: non-i.i.d. observations, contamination by noise/missing data
- **Note:** Structure learning generally harder for directed graphs (topological order unknown)

• When $(X_1, \ldots, X_p) \sim N(0, \Sigma)$, well-known fact:

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• Establishes statistical consistency of graphical Lasso (Yuan & Lin '07):

$$\widehat{\Theta} \in \arg\min_{\Theta \succeq 0} \left\{ \mathsf{trace}(\widehat{\Sigma}\Theta) - \log \mathsf{det}(\Theta) + \lambda \sum_{s \neq t} |\Theta_{st}| \right\}$$

• Only sample-based quantity is $\widehat{\Sigma}$:

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Although graphical Lasso is *penalized Gaussian MLE*, can *always* be used to estimate Θ from Σ:

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• We extend graphical Lasso to discrete-valued data (undirected case) and linear structural equation models (directed case)

• If

$$\begin{split} \|\widehat{\Sigma} - \Sigma^*\|_{\max} \precsim \sqrt{\frac{\log p}{n}} \quad \text{and} \quad \lambda \succsim \sqrt{\frac{\log p}{n}}, \\ \|\widehat{\Theta} - \Theta^*\|_{\max} \precsim \left(\sqrt{\frac{\log p}{n}} + \lambda\right) \end{split}$$

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• Deviation condition holds w.h.p. for various ensembles (e.g., sub-Gaussian)

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- Deviation condition holds w.h.p. for various ensembles (e.g., sub-Gaussian)
- Thresholding $\widehat{\Theta}$ at level $\sqrt{\frac{\log p}{n}}$ yields correct support





3 Linear structural equation models

4 Corrupted data

 (Liu et al. '09, '12): (X₁,..., X_p) follows nonparanormal distribution if (f₁(X₁),..., f_p(X_p)) ~ N(0, Σ), and f_j's monotone and differentiable

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• In general non-Gaussian setting, relationship between entries of $\Theta = \Sigma^{-1}$ and edges of G unknown

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Our results:

- Establish relationship between **augmented** inverse covariance matrices and edge structure
- New algorithms for structure learning in discrete graphs

• Binary Ising model:

$$\mathbb{P}_{\theta}(x_1,\ldots,x_{\rho}) \propto \exp\left(\sum_{s\in V} \theta_s x_s + \sum_{(s,t)\in E} \theta_{st} x_s x_t\right),$$

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 $heta \in \mathbb{R}^{p+\binom{p}{2}}, \qquad (x_1,\ldots,x_p) \in \{0,1\}^p$

• Ising models with $\theta_s=0.1, \quad \theta_{st}=2$





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• Θ is graph-structured for chain, but not loop







• However, letting $\Gamma_{\text{aug}} = \text{Cov}(X_1, X_2, X_3, X_4, \frac{X_1X_3}{2})^{-1}$ for loop:

$$\Gamma_{aug} \propto \begin{bmatrix} 115 & -2 & 109 & -2 & -114 \\ -2 & 5 & -2 & 0 & 1 \\ 109 & -2 & 114 & -2 & -114 \\ -2 & 0 & -2 & 5 & 1 \\ -114 & 1 & -114 & 1 & 119 \end{bmatrix}$$

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- **Ex:** When $U = \{1\}$, $\phi_U = (\mathbb{I}\{x_1 = 1\}, \dots, \mathbb{I}\{x_1 = m 1\})$

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- In general: Clique $C \in \mathcal{C}$ has $(m-1)^{|C|}$ indicators of nonzero states



• Triangulate G



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- Form junction tree with separator sets



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- Let $S^+ = nodes + separator sets$



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Theorem

The inverse covariance matrix of $\{\phi_U : U \in S^+\}$ from any junction tree triangulation is graph-structured: $\Gamma_{A,B} \neq 0$ iff A, B are contained in a common clique

Example: Binary Ising model



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 \bullet Statistics included in $\phi_{\mathcal{S}^+}$ depend on triangulation



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$$(\operatorname{Cov}(X_1,\ldots,X_p))^{-1}$$

• Based on exponential family representation of pdf:

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$$(\operatorname{cov}_{\theta}[\mathbb{I}(X)])^{-1} = \nabla^2 \Phi^*(\mu)$$
, where
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• Relationship between Φ^* and entropy:

$$-\Phi^*(\mu)=H(q_{ heta(\mu)}(x))=-\sum_x q_{ heta(\mu)}(x)\log q_{ heta(\mu)}(x)$$

• Junction tree theorem:

$$q(x_1,\ldots,x_p)=\frac{\prod_{C\in\mathcal{C}}q_C(x_C)}{\prod_{S\in\mathcal{S}}q_S(x_S)},$$

SO

$$H(q) = \sum_{C \in \mathcal{C}} H_C(q_C) - \sum_{S \in \mathcal{S}} H_S(q_S)$$

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 $C \in C$

• Then take Hessian

• Plug in sample covariance matrix of **augmented** vector to graphical Lasso, then compute $supp(\widehat{\Theta})$

$$\widehat{\Theta} \in \arg\min_{\Theta \succeq 0} \left\{ \mathsf{trace}(\widehat{\Sigma} \Theta) - \log \mathsf{det}(\Theta) + \lambda \sum_{s \neq t} |\Theta_{st}| \right\}$$

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• Group graphical Lasso for m > 2, similar theoretical guarantees

Problem

• However, augmented vector depends on structure of graph



• Nodewise method: recovers neighborhood N(s) for any fixed $s \in V$



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• By theorem, inverse covariance matrix over nodes and sufficient statistics of N(s) exposes neighbors of s



• By theorem, inverse covariance matrix over nodes and sufficient statistics of *N*(*s*) exposes neighbors of *s*



• Same result holds for matrix augmented by **all** *d*-subsets of $V \setminus s$

Nodewise algorithm

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- Method succeeds w.h.p. for $n \succeq 2^d \log p$
- Can incorporate noisy/missing data into Lasso-based regression

Simulations for nodewise method









4 Corrupted data

Linear structural equation models



Markov property:

 $X_j \perp \perp X_{\operatorname{Nondesc}(j)} \mid X_{\operatorname{Pa}(j)}$
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linear SEM: $X_j = b_j^T X_{\mathsf{Pa}(j)} + \epsilon_j, \qquad \epsilon_j \perp \perp X_{\mathsf{Pa}(j)}$

• $X = B^T X + \epsilon$, $X, \epsilon \in \mathbb{R}^p$ and $B \in \mathbb{R}^{p \times p}$ strictly upper triangular

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• **Goal:** Learn support of B ($B_{jk} \neq 0$ iff $j \rightarrow k$ is edge in DAG)

• Denote
$$\mathsf{Cov}(\epsilon) = \mathsf{diag}(\sigma_1^2, \dots, \sigma_p^2)$$
 and $\Theta = \mathsf{Cov}(X)^{-1}$

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Theorem

The inverse covariance matrix of X is given by

$$\Theta_{jk} = -\sigma_k^{-2} B_{jk} + \sum_{\ell > k} \sigma_\ell^{-2} B_{j\ell} B_{k\ell}, \qquad \forall j < k$$
$$\Theta_{jj} = \sigma_j^{-2} + \sum_{\ell > i} \sigma_\ell^{-2} B_{jk}^2, \qquad \forall j$$

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 $\bullet \implies \Theta_{jk} \neq 0 \text{ only when } j \rightarrow k \text{ is an edge or } j,k \text{ are parents to } \ell$

• Faithfulness assumption:

$$-\sigma_k^{-2}B_{jk} + \sum_{\ell > k} \sigma_\ell^{-2}B_{j\ell}B_{k\ell} = 0$$

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- Under faithfulness assumption, $\mathcal{M}(G) = \text{supp}(\Theta)$
- Apply graphical Lasso to estimate moralized graph

 Score-based approaches for learning DAG may be sped up with superstructure of skeleton or moralized graph (Perrier et al. '08, Ordyniak & Szeider '12)

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- Can also accommodate systematically corrupted data (next section)



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 - Missing data:

$$Z_j^{(i)} = \begin{cases} X_j^{(i)} & \text{ with prob. } 1 - \alpha \\ \star & \text{ with prob. } \alpha \end{cases}$$

• Goal: Structure learning based on $\{(Z_1^{(i)}, Z_2^{(i)}, \dots, Z_p^{(i)})\}_{i=1}^n$

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Missing data: Let

$$\widehat{Z}_{j}^{(i)} = egin{cases} rac{Z_{j}^{(i)}}{1-lpha} & ext{ if } Z_{j}^{(i)} ext{ observed} \ 0 & ext{ otherwise,} \end{cases}$$

use

$$\widehat{\boldsymbol{\Sigma}} = \frac{\widehat{Z}^T \widehat{Z}}{n} - \alpha \operatorname{diag}\left(\frac{\widehat{Z}^T \widehat{Z}}{n}\right)$$

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• Can establish deviation condition w.h.p. for modified estimators with corrupted data

Simulation study

• Graphical Lasso for dinosaur graph: probability of success for recovering 15 edges vs. rescaled sample size (with missing data)





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- Nodewise method for general discrete-valued graphs
- Modifications for corrupted data

- Computationally tractable method for structure learning in general discrete graphs
- Robustness results: Inverse covariance matrix of approximately Gaussian and/or approximately tree-structured graphs
- More general analysis of inverse covariances via exponential family representation

- P. Loh and P. Bühlmann (2013). High-dimensional learning of linear causal networks via inverse covariance estimation. ArXiv paper.
- P. Loh and M.J. Wainwright (2012). High-dimensional regression with noisy and missing data: Provable guarantees with non-convexity. *Annals of Statistics.*
- P. Loh and M.J. Wainwright (2013). Structure estimation for discrete graphical models: Generalized covariance matrices and their inverses. *Annals of Statistics.*