# CS378: Natural Language Processing Lecture 8: Expectation Maximization (EM)



Eunsol Choi



#### Recap: MEMM

Model:

$$p(y_1\dots y_n|x_1\dots x_n)=\prod_{i=1}^n p(y_i|y_1\dots y_{i-1},x_1\dots x_n)$$
 Chain rule 
$$=\prod_{i=1}^n p(y_i|y_{i-1},x_1\dots x_n)$$
 Independence assumption

Scoring:

$$p(y_i|y_{i-1},x_1...x_n) = \frac{e^{w \cdot \phi(x_1...x_n,i,y_{i-1},y_i)}}{\sum_{y'} e^{w \cdot \phi(x_1...x_n,i,y_{i-1},y')}}$$

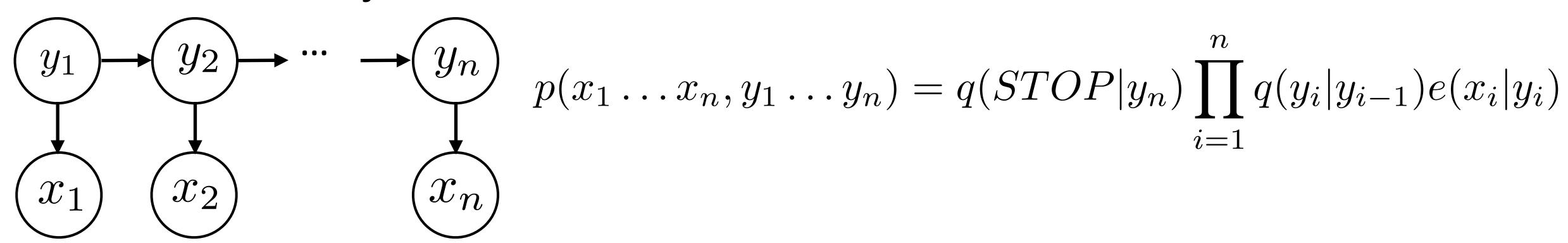
Learning Objective: log likelihood of training data

$$L = \sum_{i=1}^{n} \log P(y_i | y_1, \dots, y_{i-1}, x_1, x_2 \dots x_n)$$

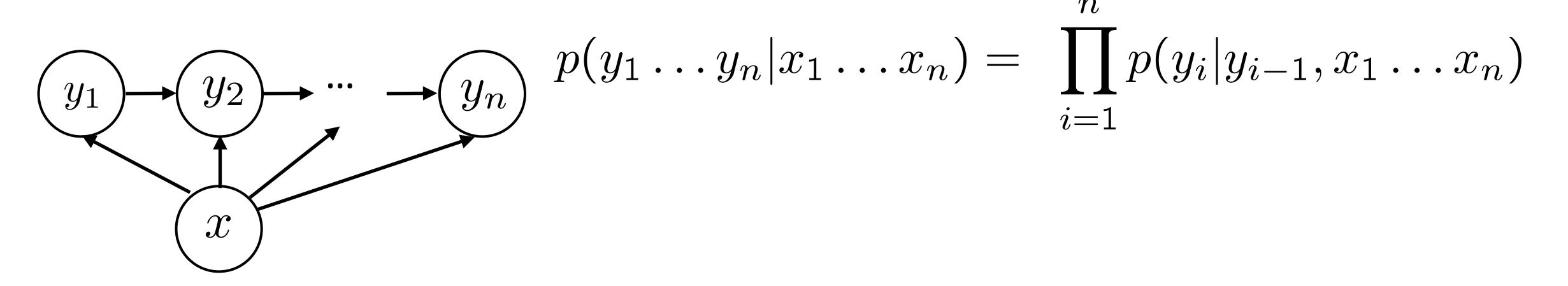


#### Recap: HMM vs. MEMM

HMM models joint distribution:



MEMM models conditional distribution:





#### The Viterbi Algorithm

Dynamic program for computing (for all i)

$$\pi(i, y_i) = \max_{y_1 \dots y_{i-1}} p(x_1 \dots x_i, y_1 \dots y_i)$$

- ▶ Iterative Computation:  $\pi(0, y_0) = \begin{cases} 1 \text{ if } y_0 == START \\ 0 \text{ otherwise} \end{cases}$
- For I = 1... n:
  - Store score

$$\pi(i, y_i) = \max_{y_{i-1}} e(x_i|y_i)q(y_i|y_{i-1})\pi(i-1, y_{i-1})$$

Store back-pointer

$$bp(i, y_i) = \arg\max_{y_{i-1}} e(x_i|y_i)q(y_i|y_{i-1})\pi(i-1, y_{i-1})$$

We end up with a single most likely sequence.



#### Recap: Decoding for MEMM

- Given your model, finding the highest scoring y
- Not very different from decoding for HMMs dynamic programming!
- Viterbi for HMMs

$$\pi(i, y_i) = \max_{y_{i-1}} e(x_i|y_i)q(y_i|y_{i-1})\pi(i-1, y_{i-1})$$

Viterbi for MEMM

$$\pi(i, y_i) = \max_{y_{i-1}} p(y_i|y_{i-1}, x_1 \dots x_n) \pi(i-1, y_{i-1})$$



#### Recap: Perceptron vs. CRF

- Both compute a score  $w \cdot \phi(\mathbf{x}, \mathbf{y})$
- Perceptron:
  - Iteratively processes the data, reacting to training errors

$$\mathbf{y}^* = \operatorname{argmax}_{\mathbf{y} \in Y} w \cdot \phi(\mathbf{x}, \mathbf{y})$$

$$w = w + \phi(\mathbf{x_i}, \mathbf{y_i}) - \phi(\mathbf{x_i}, \mathbf{y^*})$$

- Conditional Random Field:
  - Model a probability distribution over y with softmax

$$p(\mathbf{y} | \mathbf{x}) = \frac{1}{Z} \exp(\sum_{j} w \cdot \phi_{j}(\mathbf{x}, \mathbf{y}))$$

For efficient inference, both limit to local features:  $\phi(\mathbf{x}, \mathbf{y}) = \sum_{i=1}^{n} \phi(x, j, y_{j-1}, y_j)$ 



#### Another goal: Marginal Inference

- What if we are interested in the distribution of tags for each step?
- Find the marginal probability of each tag  $y_i$ :  $p(y_i | x_1, \dots x_n)$

$$p(x_1 \dots x_n, y_i) = \sum_{y_1 \dots y_{i-1}} \sum_{y_{i+1} \dots y_n} p(x_1 \dots x_n, y_1 \dots y_n)$$

$$p(x_1 \dots x_n) = \sum_{y_1 \dots y_n} p(x_1 \dots x_n, y_1 \dots y_n)$$

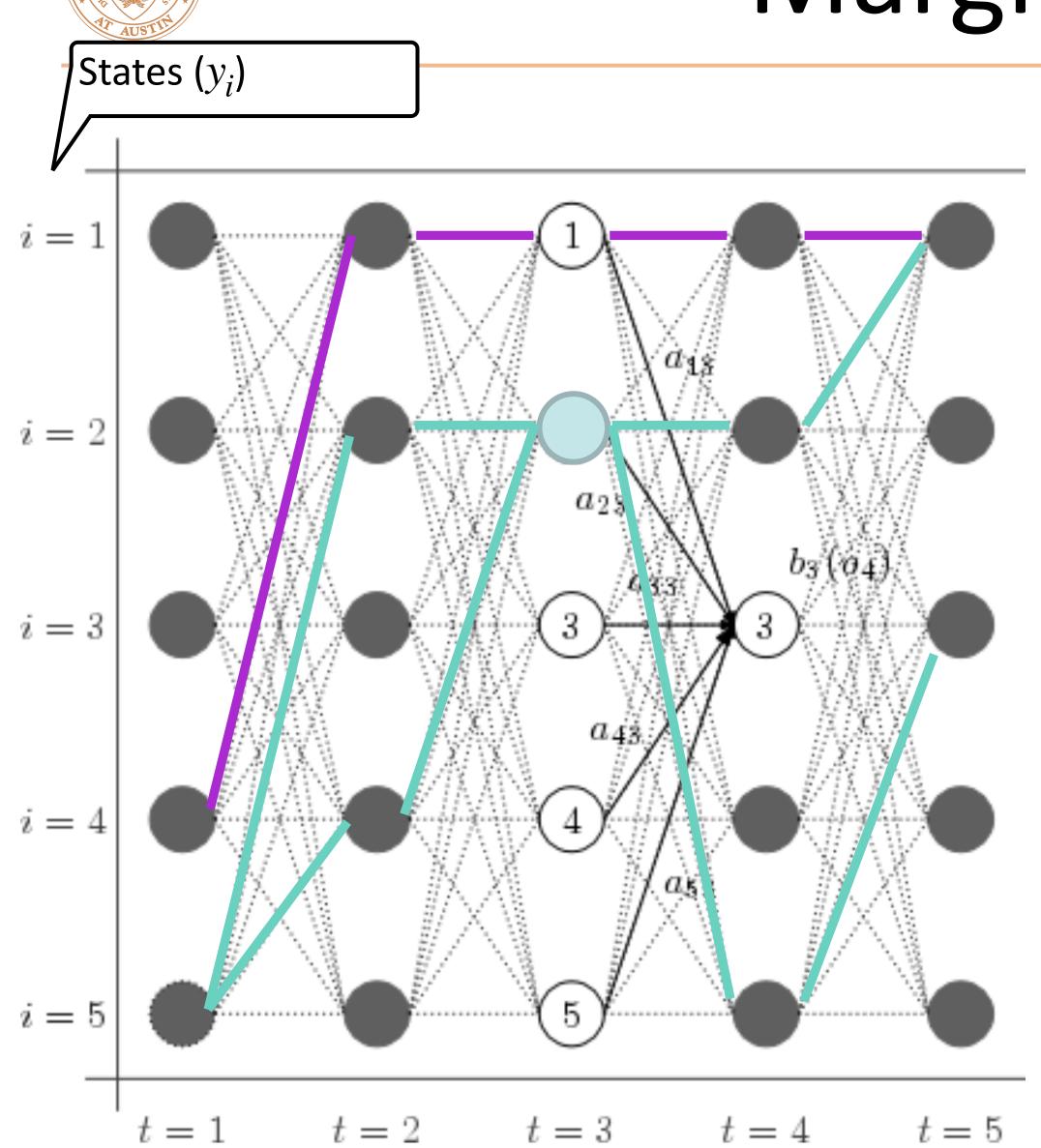
$$= \dots \dots \dots p(x_1 \dots x_n, y_i)$$

$$= \sum_{y_i} p(x_1 \dots x_n, y_i)$$

Let's compute this together!



#### Marginal Inference



$$p(y_i | x_1, \dots x_n)$$

$$P(y_3 = 2 | x_1, \dots x_n) =$$

sum of all paths through state 2 at time 3

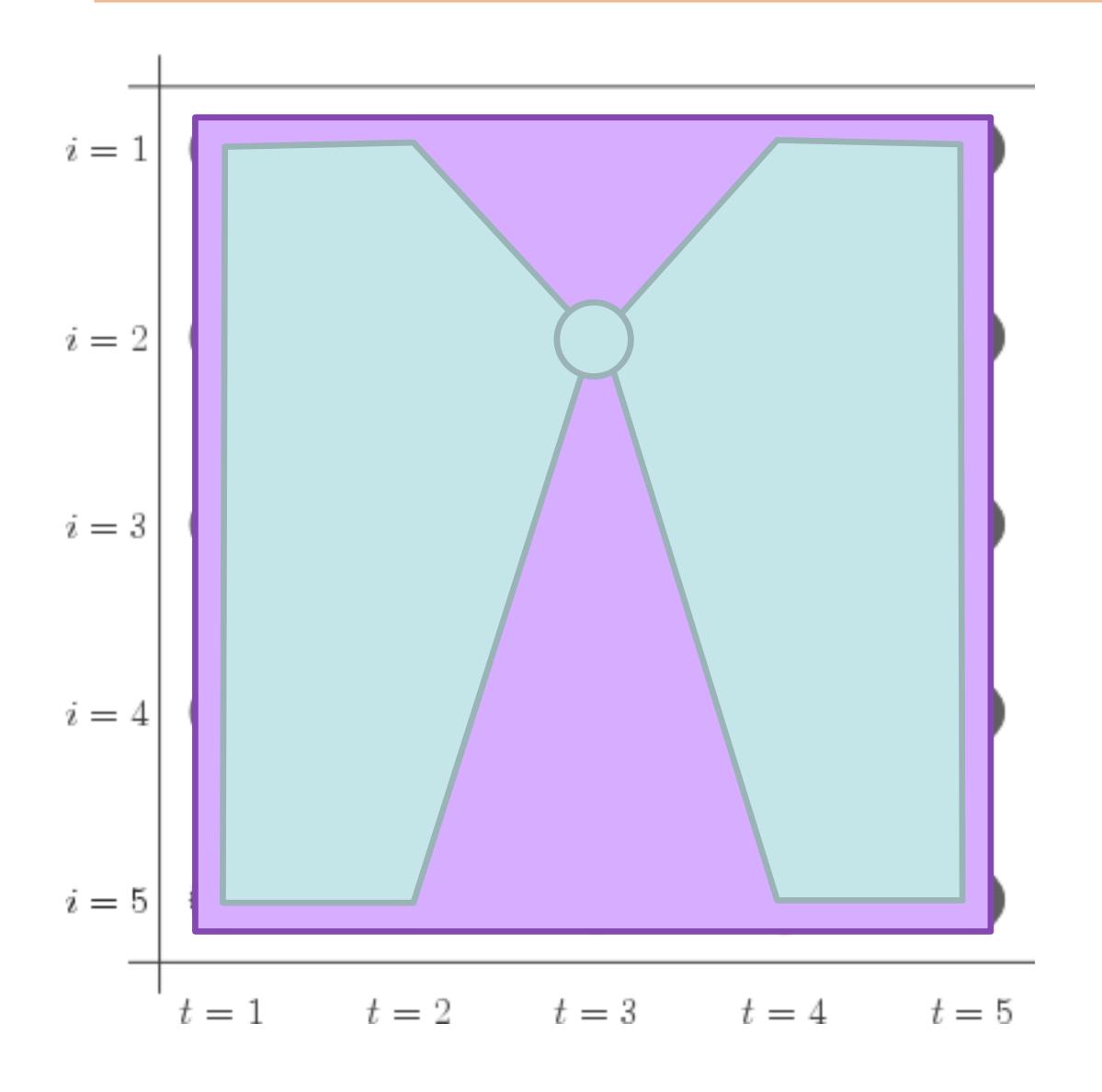
sum of all paths

The probability of input data sequence  $\;p(x_1...x_n)$ 

Time step



#### Marginal Inference



$$P(y_3 = 2 | x_1, \dots x_n) =$$

sum of all paths through state 2 at time 3 sum of all paths

Easiest and most flexible to do one pass to compute and one to compute



#### Marginal Inference

Decompose into two probabilities

$$p(x_1 \dots x_n, y_i) = p(x_1 \dots x_i, y_i) p(x_{i+1} \dots x_n | y_i)$$

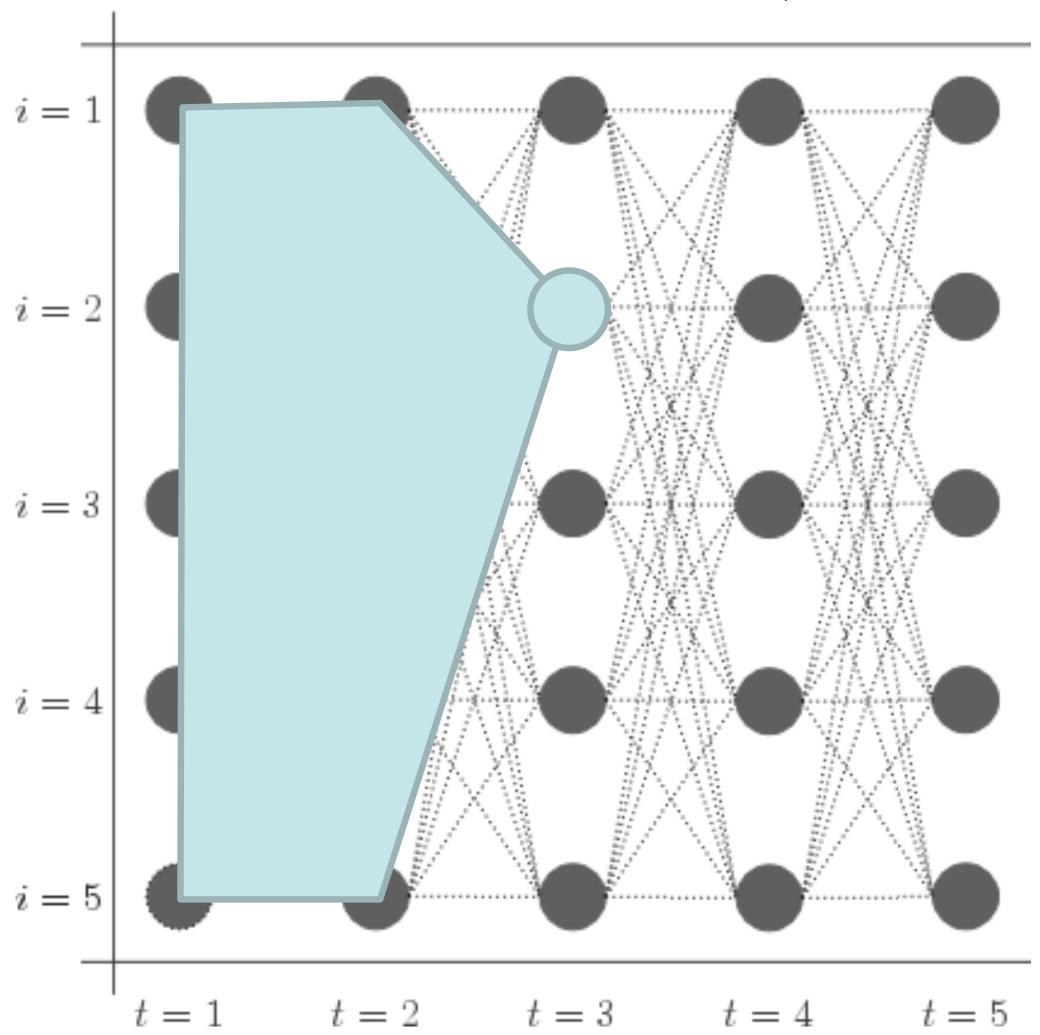
Dynamic programming on both sides

$$\alpha(i,y_i) = p(x_1 \dots x_i,y_i) = \sum_{y_1 \dots y_{i-1}} p(x_1 \dots x_i,y_1 \dots y_i)$$
 Forward pass

$$\beta(i,y_i) = p(x_{i+1}\dots x_n|y_i) = \sum_{y_{i+1}\dots y_n} p(x_{i+1}\dots x_n,y_{i+1}\dots y_n|y_i)$$
 Backward pass



$$\alpha(i, y_i) = p(x_1 \dots x_i, y_i) = \sum_{y_1 \dots y_{i-1}} p(x_1 \dots x_i, y_1 \dots y_i)$$



- Initial:  $\alpha(0, y_0) = \begin{cases} 1 \text{ if } y_0 == START \\ 0 \text{ otherwise} \end{cases}$
- Recurrence:

$$\alpha(i, y_i) = \sum_{y_{i-1}} e(x_i|y_i)q(y_i|y_{i-1})\alpha(i-1, y_{i-1})$$

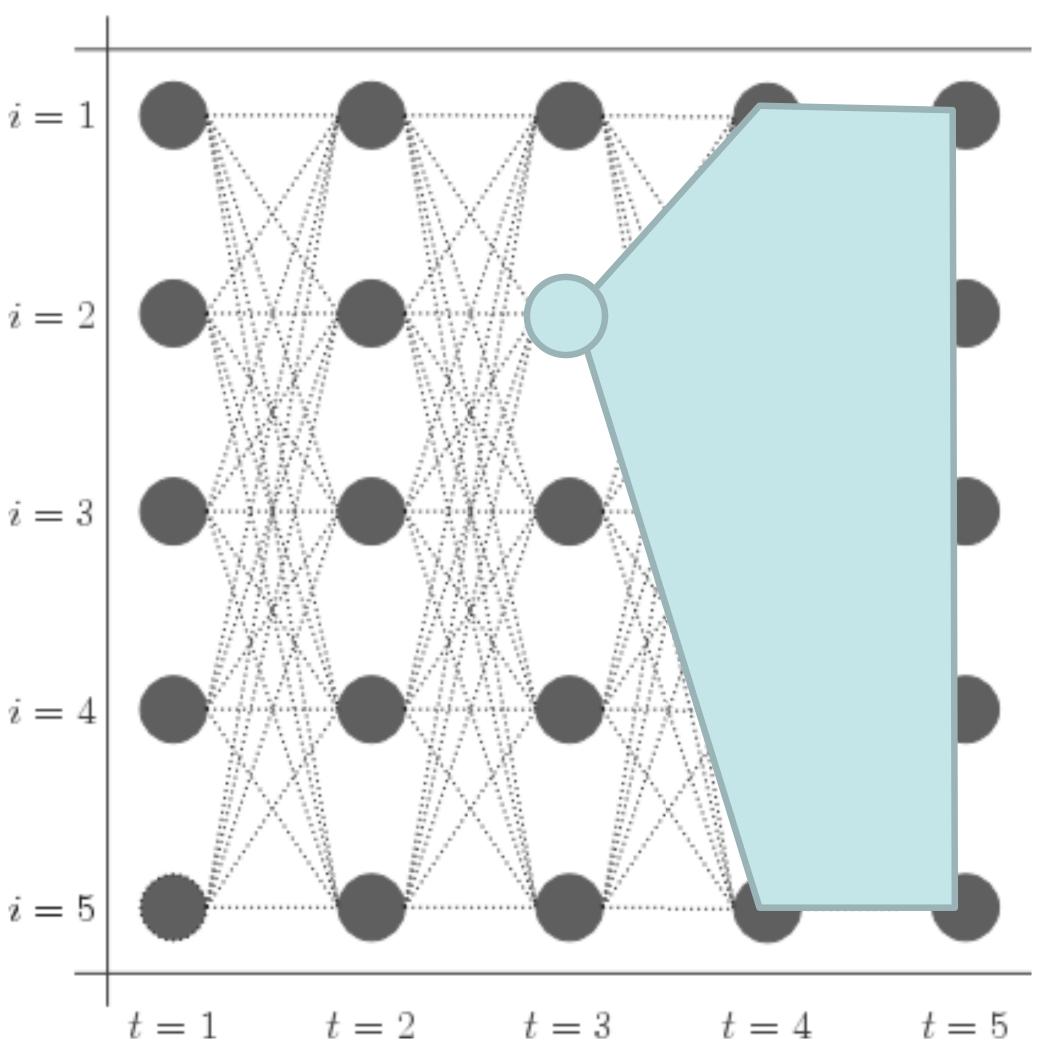
Recap: Viterbi for HMM:

$$\pi(i, y_i) = \max_{y_{i-1}} e(x_i|y_i)q(y_i|y_{i-1})\pi(i-1, y_{i-1})$$

Same as Viterbi but sum instead of max!



$$\beta(i, y_i) = p(x_{i+1} \dots x_n | y_i) = \sum_{y_{i+1} \dots y_n} p(x_{i+1} \dots x_n, y_{i+1} \dots y_n | y_i)$$



Initial:

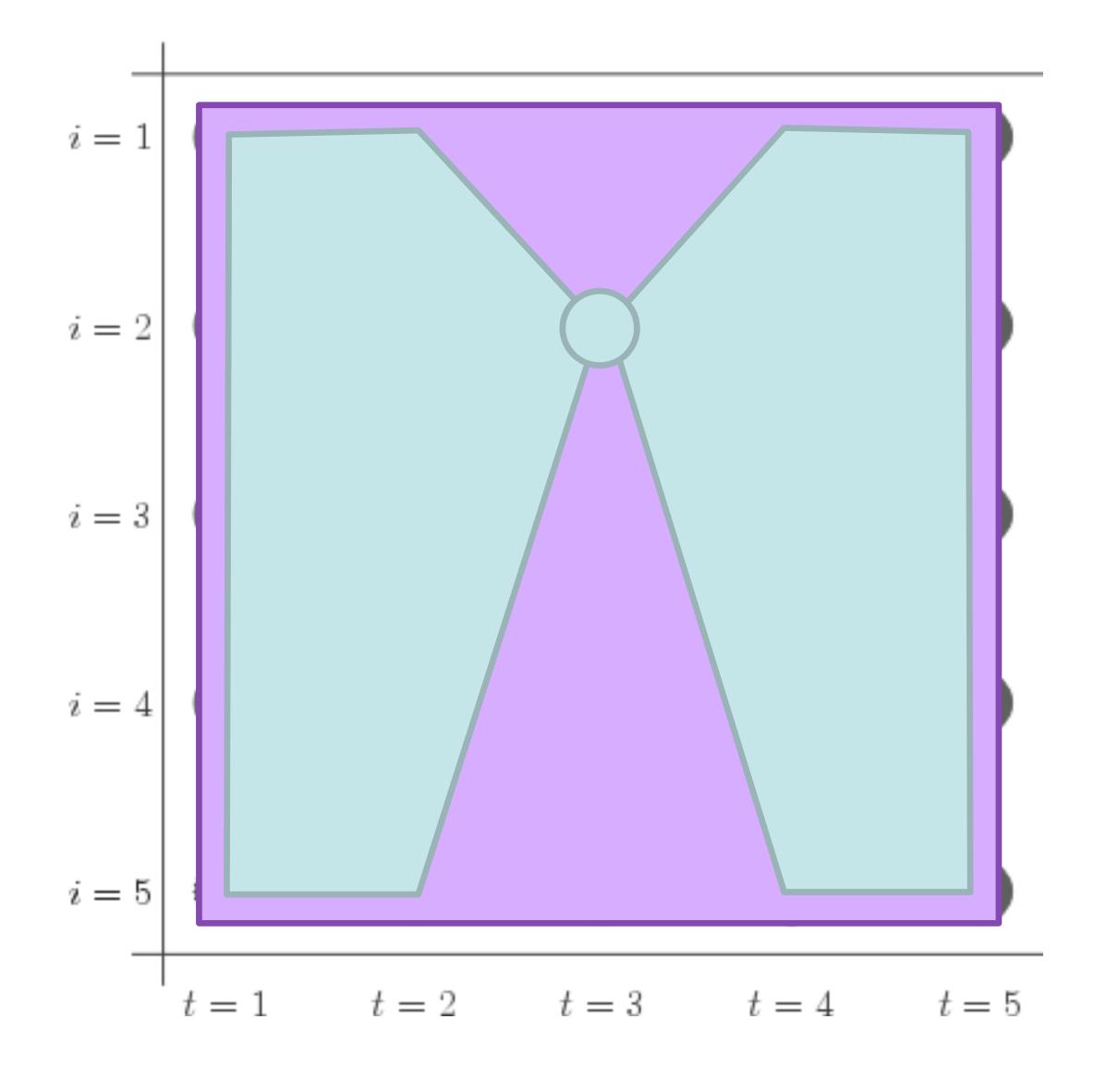
$$\beta(n, y_n) = \begin{cases} q(y_{n+1}|y_n) & \text{if } y_{n+1} = \text{STOP} \\ 0 & \text{otherwise} \end{cases}$$

Recurrence:

$$\beta(i, y_i) = \sum_{y_{i+1}} e(x_{i+1}|y_{i+1})q(y_{i+1}|y_i)\beta(i+1, y_{i+1})$$

 Big differences: count emission for the *next* timestep (not current one)





$$p(y_i|x_1...x_n) = \frac{p(x_1...x_n, y_i)}{p(x_1...x_n)}$$

$$P(y_3 = 2 \mid x_1, \dots x_n) = \frac{\alpha(3,2)\beta(3,2)}{\sum_i \alpha(3,i)\beta(3,i)}$$



#### Two Inference Methods

Viterbi algorithm

Forward Backward algorithm

Computational Costs?



#### How are these relevant in 2021?

Published as a conference paper at ICLR 2017

#### STRUCTURED ATTENTION NETWORKS

Yoon Kim\* Carl Denton\* **Luong Hoang** Alexander M. Rush

{yoonkim@seas,carldenton@college,lhoang@g,srush@seas}.harvard.edu School of Engineering and Applied Sciences Harvard University Cambridge, MA 02138, USA

#### **procedure** FORWARDBACKWARD( $\theta$ )

$$egin{aligned} & lpha[0,\langle t
angle] \leftarrow 0 \ & eta[n+1,\langle t
angle] \leftarrow 0 \ & \mathbf{for} \ i=1,\ldots,n; c \in \mathcal{C} \ \mathbf{do} \ & lpha[i,c] \leftarrow igoplus_y lpha[i-1,y] \otimes heta_{i-1,i}(y,c) \end{aligned}$$
 $\mathbf{for} \ i=n,\ldots,1; c \in \mathcal{C} \ \mathbf{do} \ & eta[i,c] \leftarrow igoplus_y eta[i+1,y] \otimes heta_{i,i+1}(c,y)$ 
 $A \leftarrow lpha[n+1,\langle t
angle]$ 

**procedure** BACKPROPFORWARDBACKWARD $(\theta, p, \nabla_p^{\mathcal{L}})$ 

 $\nabla_{\alpha}^{\mathcal{L}} \leftarrow \log p \otimes \log \nabla_{p}^{\mathcal{L}} \otimes \beta \otimes -A$  $\nabla^{\mathcal{L}}_{\beta} \leftarrow \log p \otimes \log \nabla^{\mathcal{L}}_{p} \otimes \alpha \otimes -A$ 

 $\hat{\alpha}[0,\langle t\rangle] \leftarrow$  $\hat{\beta}[n+1,\langle t]$ for i=n,. $\hat{eta}[i,c]$   $\epsilon$ for i=1,...

#### **Inside-Outside & Forward-Backward Algorithms** are just Backprop

(tutorial paper)

**Jason Eisner** 







#### **Latent Predictor Networks for Code** Generation

Wang Ling, Edward Grefenstette, Karl Moritz Andrew Senior, Fumin Wang, Phil Blunsom ACL 2016

While the number of possible paths grows exponentially,  $\alpha$  and  $\beta$  can be computed efficiently using the forward-backward algorithm for Semi-Markov models (Sarawagi and Cohen, 2005), where we associate  $P(r_t \mid y_1...y_{t-1}, x)$  to edges Hermann, Tomáš Kočiský, and  $P(s_t \mid y_1..y_{t-1}, x, r_t)$  to nodes in the Markov chain.

> The derivative  $\frac{\partial \log P(y|x)}{\partial P(s_t|y_1..y_{t-1},x,r_t)}$ can be computed using the same logic:

$$\frac{\partial \alpha_{t,s_{t}} P(s_{t} \mid y_{1}..y_{t-1}, x, r_{t}) \beta_{t+|s_{t}|-1} + \xi_{r_{t}}}{P(y \mid x) \partial P(s_{t} \mid y_{1}..y_{t-1}, x, r_{t})} = \frac{\alpha_{t,r_{t}} \beta_{t+|s_{t}|-1}}{\alpha_{|y|+1}}$$

"The inside-outside algorithm is the hardest algorithm I know."



- a senior NLP researcher, in the 1990's



# Can we learn the latent states without supervised training dataset?



#### Sequence Labeling: Partially Observed

Model parameters

- We have a sequence  ${\bf x}$  and  ${\bf y}$ , and a joint distribution  $p({\bf x},{\bf y}\,|\,\theta)$
- So far, we had fully observable data  $(\mathbf{x}_i, \mathbf{y}_i)$  pairs,

Learning Objective 
$$L(\theta) = \sum_{i=1}^{n} \log P(\mathbf{x}_i, \mathbf{y}_i | \theta)$$

What if we only have access to observations x?

$$L(\theta) = \sum_{i} \log P(\mathbf{x_i} | \theta)$$



#### Maximum Likelihood Estimate

- Data points  $x_1, x_2, x_3 \dots x_n$
- ightharpoonup Parameter vector  $\, heta$  and a parameter space  $\,\Omega\,$
- Probability distribution  $P(x | \theta)$  any  $\theta$  in  $\Omega$ .

Likelihood 
$$P(x_1, x_2, \dots, x_n | \theta) = \prod_{i=1}^n P(x_i | \theta)$$
  
Log likelihood:  $L(\theta) = \sum_{i=1}^n \log P(x_i | \theta)$ 

Goal: increase the probability of training data!



#### Expectation Maximization

Learning objective:

$$L(\theta) = \sum_{i} \log \sum_{y \in \mathcal{Y}} P(x_i, y \mid \theta)$$

The EM (Expectation Maximization) algorithm is a method for finding

$$\theta_{MLE} = \arg\max_{\theta} L(\theta) = \arg\max_{\theta} \sum_{i} \log \sum_{y \in \mathcal{Y}} P(x_i, y \mid \theta)$$

We will look into EM in HMM!

$$p(x_1 \dots x_n, y_1 \dots y_n) = q(STOP|y_n) \prod_{i=1}^n q(y_i|y_{i-1})e(x_i|y_i)$$



#### General Idea

- Initially guess the model parameters  $\theta$ 

Iterate two steps:

 Expectation step: Use the current parameters (and observations) to reconstruct hidden states

 Maximization step: Use that hidden states (and observations) to reestimate the parameters



#### EM for HMM

Maximum Likelihood Parameters (supervised):

$$q_{ML}(y_i|y_{i-1}) = \frac{c(y_{i-1}, y_i)}{c(y_{i-1})} \qquad e_{ML}(x|y) = \frac{c(y, x)}{c(y)}$$

- For unsupervised learning, replace actual count with expected counts
- Expected emission counts:

(expected) count(NN 
$$\rightarrow$$
 apple) =  $\sum_{i} p(y_i = \text{NN}, x_i = \text{apple}|x_1...x_n)$   
=  $\sum_{i:x_i = \text{apple}} p(y_i = \text{NN}|x_1...x_n)$ 



#### EM Intuition

What we want is...

$$p(y_i|x_1...x_n) = \frac{p(x_1...x_n, y_i)}{p(x_1...x_n)}$$

We can compute:

(expected) count(NN) = 
$$\sum_{i} p(y_i = \text{NN}|x_1...x_n)$$

If we have....

$$p(y_i y_{i+1} | x_1 ... x_n) = \frac{p(x_1 ... x_n, y_i, y_{i+1})}{p(x_1 ... x_n)}$$

Then we can compute expected transition counts:

(expected) count(NN 
$$\rightarrow$$
 VB) =  $\sum_{i} p(y_i = \text{NN}, y_{i+1} = \text{VB}|x_1...x_n)$ 

Above marginals can be computed from followings:

$$p(x_1...x_n, y_i) = \alpha(i, y_i)\beta(i, y_i)$$
  

$$p(x_1...x_n, y_i, y_{i+1}) = \alpha(i, y_i)q(y_{i+1}|y_i)e(x_{i+1}|y_{i+1})\beta(i+1, y_{i+1})$$



#### Expectation Maximization

- Initialize transition and emission parameters:
  - random, uniform, or more informed initialization
- Iterate until convergence
  - ► E-step: computing expected counts

unts (expected) count(NN) = 
$$\sum_{i} p(y_i = \text{NN} | x_1...x_n)$$
  
(expected) count(NN  $\rightarrow$  VB) =  $\sum_{i} p(y_i = \text{NN}, y_{i+1} = \text{VB} | x_1...x_n)$   
(expected) count(NN  $\rightarrow$  apple) =  $\sum_{i} p(y_i = \text{NN}, x_i = \text{apple} | x_1...x_n)$ 

M-step: computing new transition and emission parameters

$$q_{ML}(y_i|y_{i-1}) = \frac{c(y_{i-1}, y_i)}{c(y_{i-1})} \qquad e_{ML}(x|y) = \frac{c(y, x)}{c(y)}$$

Convergence? Yes. Global Optimum? No.

**function** FORWARD-BACKWARD(observations of len T, output vocabulary V, hidden state set Q) **returns** HMM=(A,B)

initialize A and B iterate until convergence

E-step

$$\gamma_t(j) = \frac{\alpha_t(j)\beta_t(j)}{P(O|\lambda)} \,\forall t \text{ and } j$$

$$\xi_t(i,j) = \frac{\alpha_t(i)a_{ij}b_j(o_{t+1})\beta_{t+1}(j)}{\alpha_T(N)} \,\forall t, i, \text{ and } j$$

M-step

$$\hat{a}_{ij} = \frac{\sum_{t=1}^{T-1} \xi_t(i,j)}{\sum_{t=1}^{T-1} \sum_{j=1}^{N} \xi_t(i,j)}$$

$$\hat{b}_j(v_k) = \frac{\sum_{t=1}^{T} \sum_{j=1}^{T} \gamma_t(j)}{\sum_{t=1}^{T} \gamma_t(j)}$$

Equivalent to the procedure given in the textbook (J&M Appendix A) – slightly different notations

return A, B



- You are studying global warming. You cannot find records of weather, but you can find records of how much ice cream was consumed each day. Can you estimate the weather history from the ice cream history?
  - Observations (x): Number of ice cream purchase
    - **1**, 2, 3
  - State (y): Weather
    - {C (cold), H (hot)}



If today is cold (C) or hot (H), how many cones did I prob. eat?

	P( C)	P( H)
P(1 )	0.7	0.1
P(2 )	0.2	0.2
P(3 )	0.1	0.7

	P( C)	P( H)	P( start)
P(C )	0.8	0.1	0.5
P(H )	0.1	8.0	0.5
P(Stop )	0.1	0.1	0

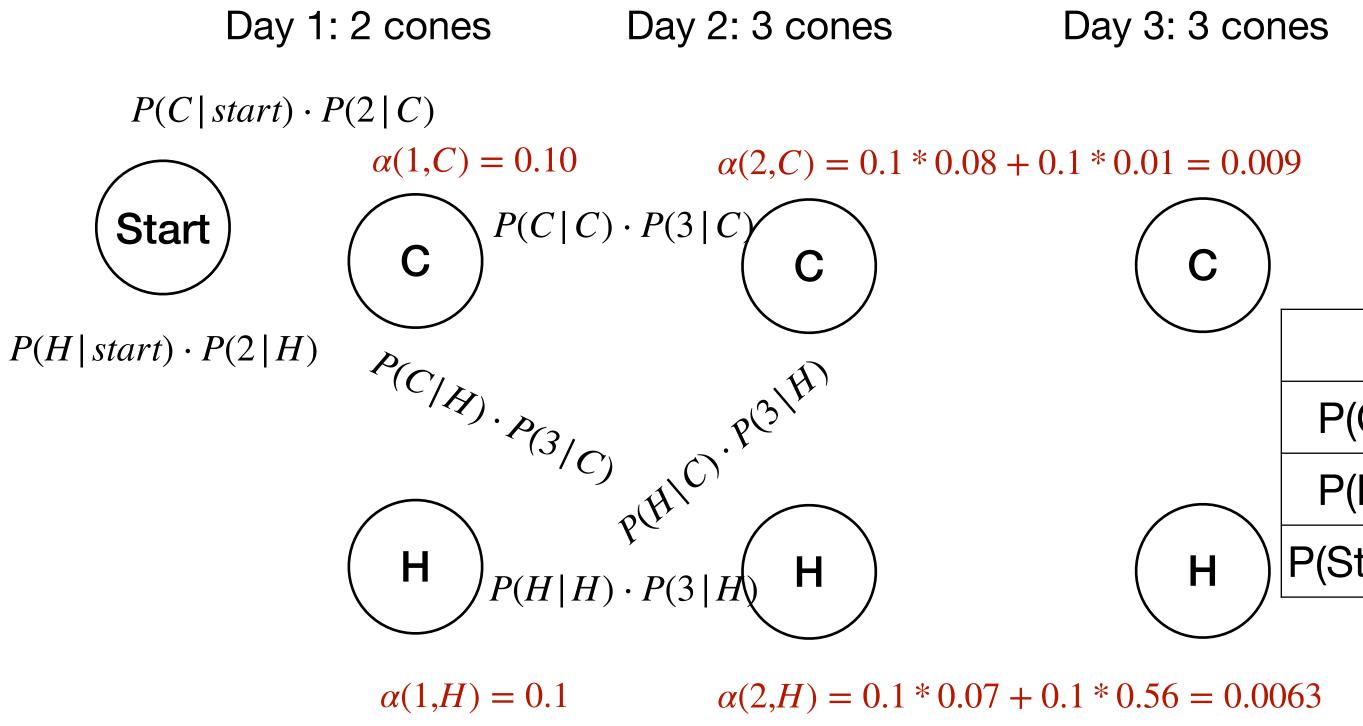
If today is cold (C) or hot (H), what will tomorrow's weather be?

Maximum Likelihood Parameters (supervised):

$$e_{ML}(x|y) = \frac{c(y,x)}{c(y)}$$
 
$$q_{ML}(y_i|y_{i-1}) = \frac{c(y_{i-1},y_i)}{c(y_{i-1})}$$

 Now, we do not have the weather record (real counts), so we start with guessed probability table, and compute expected counts





	P( C)	P( H)
P(1 )	0.7	0.1
P(2 )	0.2	0.2
P(3 )	0.1	0.7

	P( C)	P( H)	P( start)
P(C )	8.0	0.1	0.5
P(H )	0.1	0.8	0.5
P(Stop )	0.1	0.1	0

•	We compute f	forward	and	bac	kward	, from	whicl	n we
	compute P(y_	$i x$ ) at $\epsilon$	each	day	(i).			

$$p(x_1...x_n, y_i) = \alpha(i, y_i)\beta(i, y_i)$$

	Ice		
Day#	Creams	p(□C)	p(□H)
	1 2	0.129	0.871
	2 3	0.023	0.977
	3 3	0.011	0.989
	4 2	0.027	0.973
	5 3	0.013	0.987
	6 <b>2</b>	0.032	0.968
	7 3	0.022	0.978
	8 <b>2</b>	0.069	0.931
	9 2	0.089	0.911
1	.0 3	0.082	0.918
1	1 1	0.248	0.752
	.2 3	0.144	0.856
1	.3 3	0.221	0.779
1	4 1	0.887	0.113
1	.5 1	0.98	0.02
	.6 1	0.991	0.009
	.7 <b>2</b>	0.977	0.023
	.8 1	0.994	0.006
	9 1	0.994	0.006
	0 1	0.977	0.023
	1 3	0.857	0.143
	2 1	0.962	0.038
	3 2	0.961	0.039
	4 1	0.989	0.011
	5 1	0.985	0.015
	6 1	0.926	0.074
	7 2	0.507	0.493
	8 3	0.087	0.913
	9 3	0.032	0.968
	0 2	0.053	0.947
	1 3	0.045	0.955
	2 2	0.146	0.854
	3 2	0.225	0.775
3	ai Z	14.679	18.321



p(□C)	p(□H)	p(□C,1)	p(□C,2)	p(□C,3)	p(□H,1)	p(□H,2)	p(□H,3)	p(C□C)	p(H□C)	p(C□H)	p(H□H)
0.129	0.871	0	0.129	0	0	0.871	0	#N/A	#N/A	#N/A	#N/A
0.023	0.977	0	0	0.023	0	0	0.977	0.021	0.003	0.109	0.868
0.011	0.989	0	0	0.011	0	0	0.989	0.006	0.005	0.017	0.972
0.027	0.973	0	0.027	0	0	0.973	0	0.006	0.021	0.005	0.969
0.013	0.987	0	0	0.013	0	0	0.987	0.007	0.005	0.02	0.968
0.032	0.968	0	0.032	0	0	0.968	0	800.0	0.024	0.005	0.963
0.022	0.978	0	0	0.022	0	0	0.978	0.012	0.009	0.02	0.959
0.069	0.931	0	0.069	0	0	0.931	0	0.017	0.052	0.005	0.927
0.089	0.911	0	0.089	0	0	0.911	0	0.05	0.038	0.018	0.893
0.082	0.918	0	0	0.082	0	0	0.918	0.057	0.025	0.031	0.886
0.248	0.752	0.248	0	0	0.752	0	0	0.077	0.171	0.005	0.747
0.144	0.856	0	0	0.144	0	0	0.856	0.131	0.013	0.117	0.739
0.221	0.779	0	0	0.221	0	0	0.779	0.128	0.093	0.016	0.762
0.887	0.113	0.887	0	0	0.113	0	0	0.221	0.666	0.001	0.113
0.98	0.02	0.98	0	0	0.02	0	0	0.884	0.095	0.003	0.018
0.991	0.009	0.991	0	0	0.009	0	0	0.975	0.016	0.005	0.005
0.977	0.023	0	0.977	0	0	0.023	0	0.973	0.004	0.018	0.005
0.994	0.006	0.994	0	0	0.006	0	0	0.974	0.019	0.003	0.004
0.994	0.006	0.994	0	0	0.006	0	0	0.989	0.005	0.005	0.002
0.977	0.023	0.977	0	0	0.023	0	0	0.974	0.003	0.019	0.004
0.857	0.143	0	0	0.857	0	0	0.143	0.855	0.002	0.122	0.021
0.962	0.038	0.962	0	0	0.038	0	0	0.853	0.109	0.004	0.034
0.961	0.039	0	0.961	0	0	0.039	0	0.944	0.017	0.018	0.021
0.989	0.011	0.989	0	0	0.011	0	0	0.957	0.032	0.004	0.008
0.985	0.015	0.985	0	0	0.015	0	0	0.978	0.007	0.011	0.005
0.926	0.074	0.926	0	0	0.074	0	0	0.924	0.003	0.061	0.012
0.507	0.493	0	0.507	0	0	0.493	0	0.505	0.001	0.421	0.072
0.087	0.913	0	0	0.087	0	0	0.913	0.085	0.002	0.421	0.492
0.032	0.968	0	0	0.032	0	0	0.968	0.026	0.006	0.061	0.907
0.053	0.947	0	0.053	0	0	0.947	0	0.022	0.031	0.01	0.937
0.045	0.955	0	0	0.045	0	0	0.955	0.028	0.017	0.025	0.931
0.146	0.854	0	0.146	0	0	0.854	0	0.04	0.106	0.005	0.849
0.225	0.775	0	0.225	0	0	0.775	0	0.13	0.095	0.016	0.759
14.679	18.321	9.931	3.212	1.537	1.069	7.788	9.463	12.855	1.695	1.599	15.85

 With the expected counts for hot and cold days for each day, we compute the following

$$p(x_1...x_n, y_i) = \alpha(i, y_i)\beta(i, y_i)$$
  

$$p(x_1...x_n, y_i, y_{i+1}) = \alpha(i, y_i)q(y_{i+1}|y_i)e(x_{i+1}|y_{i+1})\beta(i+1, y_{i+1})$$

 Use these values that count to re-compute transition probability and emission probability

	P( C)	P( H)
P(1 )	0.6765	0.0584
P(2 )	0.2188	0.4251
P(3 )	0.1047	0.5165

	P( C)	P( H)	P( start)
P(C )	0.8757	0.0925	0.1291
P(H )	0.109	0.8652	0.8709
P(Stop )	0.0153	0.0423	0



# Quiz: p(S1) vs. p(S2)

- S1 = Colorless green ideas sleep furiously.
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- How would p(S1) and p(S2) compare based on (smoothed) bigram language models?
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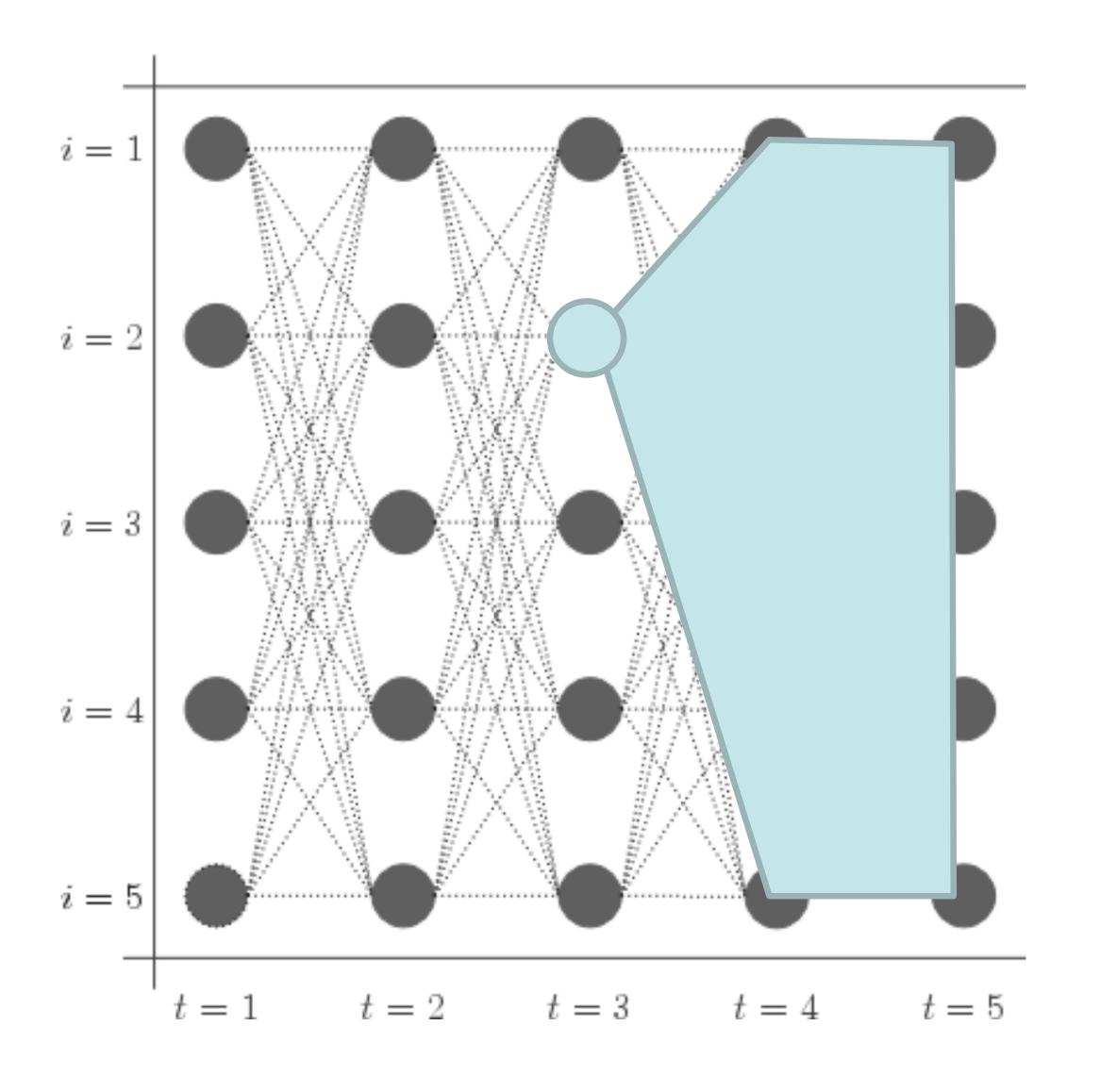
# Summary: Sequence Models

- Sequence Modeling Problems in NLP
- Generative Model: Hidden Markov Models (HMM)

Discriminative Model:
 Maximum Entropy Markov Models (MEMM)
 Conditional Random Fields

Unsupervised Learning: Expectation Maximization





Initial:

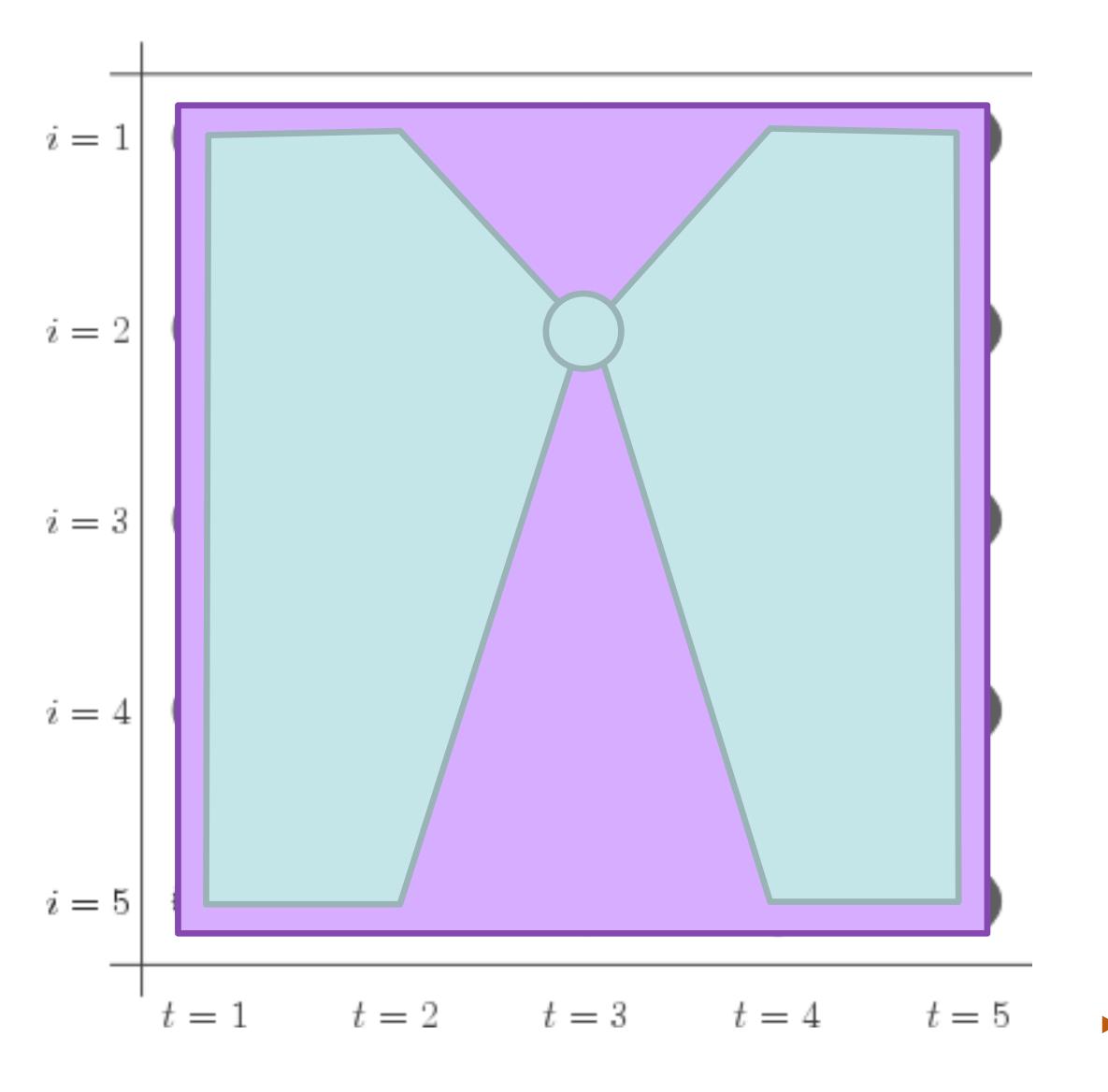
$$\beta_n(s) = 1$$

Recurrence:

$$\beta_t(s_t) = \sum_{s_{t+1}} \beta_{t+1}(s_{t+1}) \exp(\phi_e(s_{t+1}, t+1, \mathbf{x}))$$
$$\exp(\phi_t(s_t, s_{t+1}))$$

 Big differences: count emission for the *next* timestep (not current one)





$$\alpha_1(s) = \exp(\phi_e(s, 1, \mathbf{x}))$$

$$\alpha_t(s_t) = \sum_{s_{t-1}} \alpha_{t-1}(s_{t-1}) \exp(\phi_e(s_t, t, \mathbf{x}))$$

$$\exp(\phi_t(s_{t-1}, s_t))$$

$$\beta_n(s) = 1$$

$$\beta_t(s_t) = \sum_{s_{t+1}} \beta_{t+1}(s_{t+1}) \exp(\phi_e(s_{t+1}, t+1, \mathbf{x}))$$
$$\exp(\phi_t(s_t, s_{t+1}))$$

$$P(s_3 = 2|\mathbf{x}) = \frac{\alpha_3(2)\beta_3(2)}{\sum_i \alpha_3(i)\beta_3(i)}$$

• What is the denominator here?  $P(\mathbf{x})$ 



#### Forward Backward for HMM

- Two passes: one forward, one back
- Forward:

$$\alpha(0, y_0) = \begin{cases} 1 \text{ if } y_0 == START \\ 0 \text{ otherwise} \end{cases}$$

• For i = 1 ... n

$$\alpha(i, y_i) = \sum_{y_{i-1}} e(x_i|y_i)q(y_i|y_{i-1})\alpha(i-1, y_{i-1})$$

Backward:

$$\beta(n, y_n) = \begin{cases} q(y_{n+1}|y_n) & \text{if } y_{n+1} = \text{STOP} \\ 0 & \text{otherwise} \end{cases}$$

► For i = n-1 ... 0

$$\beta(i, y_i) = \sum_{y_{i+1}} e(x_{i+1}|y_{i+1})q(y_{i+1}|y_i)\beta(i+1, y_{i+1})$$

$$p(x_1...x_n, y_i) = \alpha(i, y_i)\beta(i, y_i)$$



# Other Marginal Inference

Can we compute this?

$$p(x_1...x_n) = \sum_{y_i} p(x_1...x_n, y_i)$$

Relation with forward quantity?

$$\alpha(i, y_i) = p(x_1 \dots x_i, y_i) = \sum_{y_1 \dots y_{i-1}} p(x_1 \dots x_i, y_1 \dots y_i)$$

$$p(x_1 \dots x_n) = \sum_{y_1 \dots y_n} p(x_1 \dots x_n, y_1 \dots y_{n+1})$$

$$= \dots \dots \alpha(n, y_n)$$

$$= \sum_{y_n} q(STOP|y_n)\alpha(n, y_n) := \alpha(n+1, STOP)$$



#### EM Intuition

What we want is...

$$p(y_i|x_1...x_n) = \frac{p(x_1...x_n, y_i)}{p(x_1...x_n)}$$

We can compute:

(expected) count(NN) = 
$$\sum_{i} p(y_i = \text{NN}|x_1...x_n)$$

If we have....

$$p(y_i y_{i+1} | x_1 ... x_n) = \frac{p(x_1 ... x_n, y_i, y_{i+1})}{p(x_1 ... x_n)}$$

Then we can compute expected transition counts:

(expected) count(NN 
$$\rightarrow$$
 VB) =  $\sum_{i} p(y_i = \text{NN}, y_{i+1} = \text{VB}|x_1...x_n)$ 

Above marginals can be computed from followings:

$$p(x_1...x_n, y_i) = \alpha(i, y_i)\beta(i, y_i)$$
  

$$p(x_1...x_n, y_i, y_{i+1}) = \alpha(i, y_i)q(y_{i+1}|y_i)e(x_{i+1}|y_{i+1})\beta(i+1, y_{i+1})$$



#### Expectation Maximization

- Initialize transition and emission parameters:
  - random, uniform, or more informed initialization
- Iterate until convergence
  - ► E-step: computing expected counts

ounts (expected) count(NN) = 
$$\sum_{i} p(y_i = \text{NN} | x_1...x_n)$$
  
(expected) count(NN  $\rightarrow$  VB) =  $\sum_{i} p(y_i = \text{NN}, y_{i+1} = \text{VB} | x_1...x_n)$   
(expected) count(NN  $\rightarrow$  apple) =  $\sum_{i} p(y_i = \text{NN}, x_i = \text{apple} | x_1...x_n)$ 

M-step: computing new transition and emission parameters

$$q_{ML}(y_i|y_{i-1}) = \frac{c(y_{i-1}, y_i)}{c(y_{i-1})} \qquad e_{ML}(x|y) = \frac{c(y, x)}{c(y)}$$

Convergence? Yes. Global Optimum? No.

**function** FORWARD-BACKWARD(observations of len T, output vocabulary V, hidden state set Q) **returns** HMM=(A,B)

initialize A and B iterate until convergence

E-step

$$\gamma_t(j) = \frac{\alpha_t(j)\beta_t(j)}{P(O|\lambda)} \,\forall t \text{ and } j$$

$$\xi_t(i,j) = \frac{\alpha_t(i)a_{ij}b_j(o_{t+1})\beta_{t+1}(j)}{\alpha_T(N)} \,\forall t, i, \text{ and } j$$

M-step

$$\hat{a}_{ij} = \frac{\sum_{t=1}^{T-1} \xi_t(i,j)}{\sum_{t=1}^{T-1} \sum_{j=1}^{N} \xi_t(i,j)}$$

$$\hat{b}_j(v_k) = \frac{\sum_{t=1}^{T} \sum_{j=1}^{T} \gamma_t(j)}{\sum_{t=1}^{T} \gamma_t(j)}$$

Equivalent to the procedure given in the textbook (J&M Appendix A) – slightly different notations

return A, B



#### How is this even possible?

- I water the garden everyday
- Saw a weird bug in that garden ...
- While I was thinking of an equation ...

#### Noun

S: (n) garden (a plot of ground where plants are cultivated)

<u>S:</u> (n) **garden** (the flowers or vegetables or fruits or herbs that are cultivated in a garden)

S: (n) garden (a yard or lawn adjoining a house)

#### Verb

S: (v) garden (work in the garden) "My hobby is gardening"

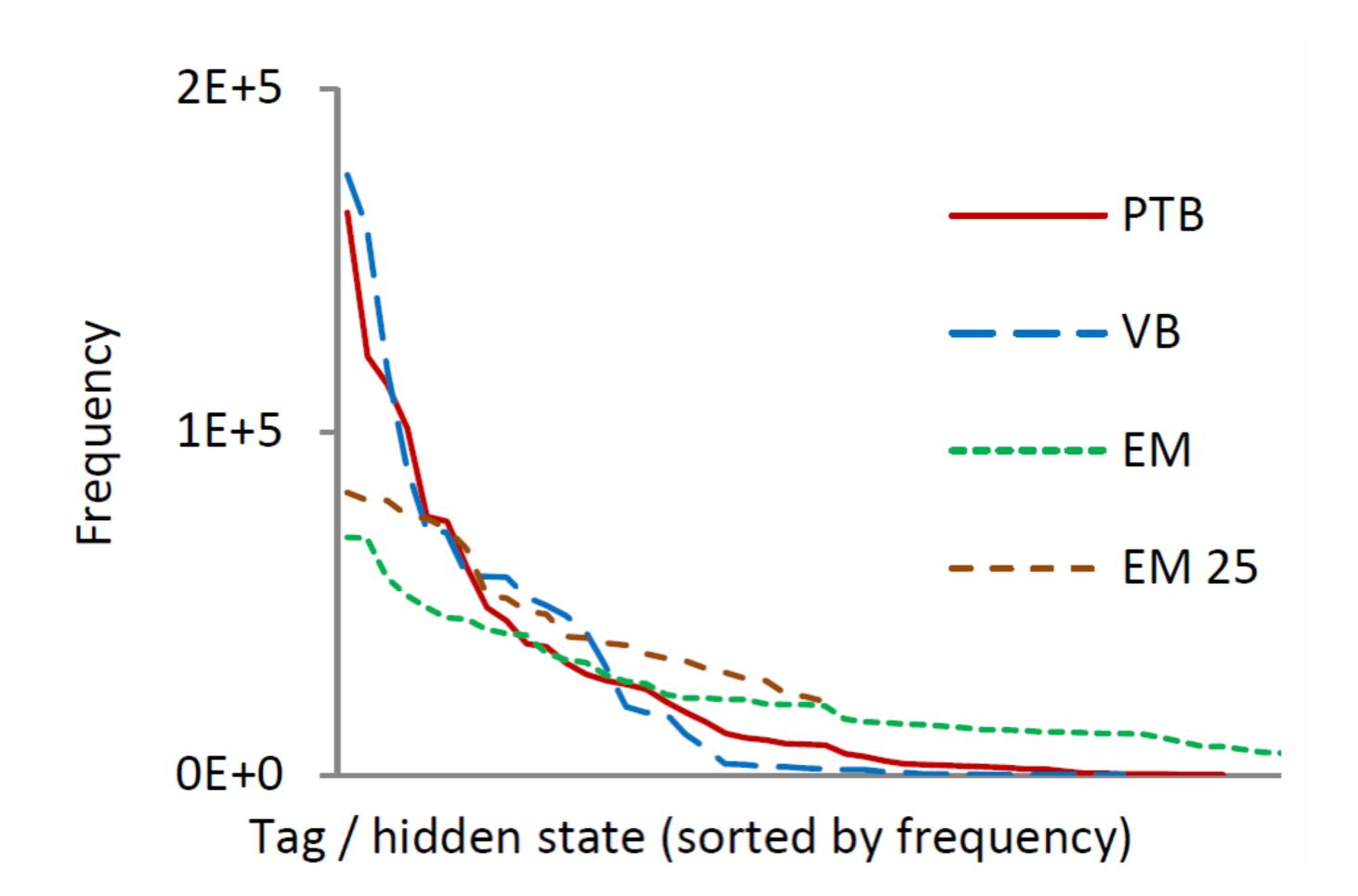
#### **Adjective**

S: (adj) garden (the usual or familiar type) "it is a common or garden sparrow"

We can start with a tag dictionary (a list of potential POSs per word)



#### Does EM learn good POS tagger?



HMMs estimated by EM generally assign a roughly equal number of word tokens to each hidden state, while the empirical distribution of tokens to POS tags is highly skewed



#### HMM vs. CRF

$$P(y_i = s | \mathbf{x}) = \frac{\text{forward}_i(s) \text{backward}_i(s)}{\sum_{s'} \text{forward}_i(s') \text{backward}_i(s')}$$

Posterior is derived from the parameters and the data (conditioned on x!)

$$P(x_i|y_i), P(y_i|y_{i-1})$$

 $P(y_i|\mathbf{x}), P(y_{i-1}, y_i|\mathbf{x})$ 

**HMM** 

Model parameter (usually multinomial distribution)

Inferred quantity from forward-backward

**CRF** 

Undefined (model is by definition conditioned on x)

Inferred quantity from forward-backward



#### POS Results

Baseline: assign each word its most frequent tag: ~90% accuracy

Trigram HMM: ~95% accuracy / 55% on unknown words

► TnT tagger (Brants 1998, tuned HMM): 96.2% accuracy / 86.0% on unks

MEMM [Ratnaparkhi 1996]: 96.8% accuracy / 86.9% on dunks

• EM for HMM: 74.7%



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#### Summary: Sequence Models

- Generative vs. Discriminative
- Structured or not
  - Independent predictions are effective, but global structure helps
  - But need to be careful about computational overhead
- Model expressivity
  - Expressive feature set can be better
  - But more expensive, overfitting
- The higher accuracy of discriminative models comes at the price of much slower training